

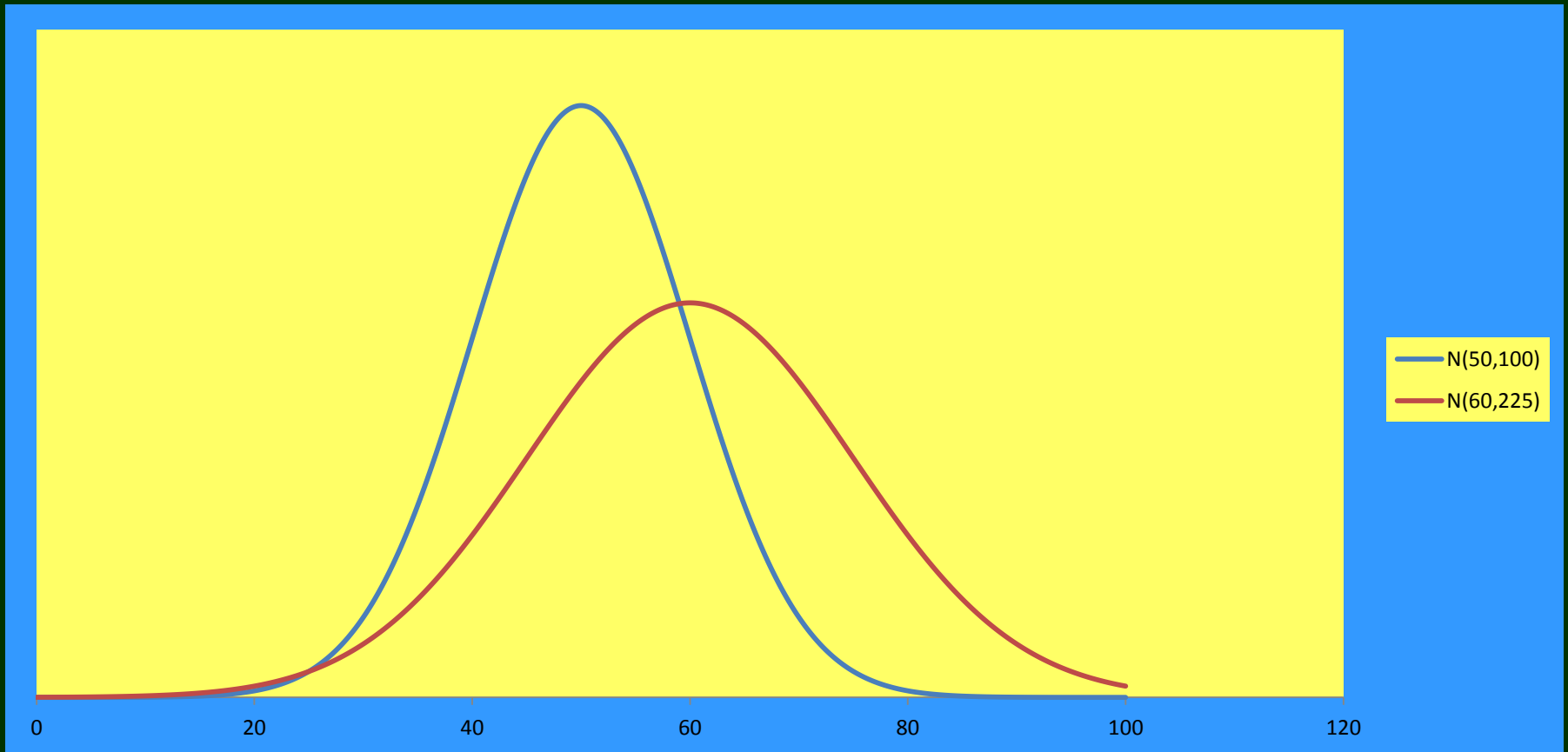
Satterthwaite's Approximation for Degrees of Freedom

Art Instruction Effect on Reading Development

J.C. Mills (1973), "The Effect of Art Instruction Upon a Reading Development Test: An Experimental Study with Rural Appalachian Children," *Studies in Art Education*, Vol. 14, #3, pp.4-8

Setting

- Comparing Means from 2 Normal Distributions
- Small Samples (Computer Packages Solve for any sizes)
- Distributions have Possibly Different Variances



Case 1 – Variances are Equal

$$Y_{11}, \dots, Y_{1n_1} \sim N(\mu_1, \sigma^2) \quad Y_{21}, \dots, Y_{2n_2} \sim N(\mu_2, \sigma^2) \quad \sigma_1^2 = \sigma_2^2 = \sigma^2 \quad \{Y_1\} \perp \{Y_2\}$$

$$\bar{Y}_i = \frac{\sum_{j=1}^{n_i} Y_{ij}}{n_i} \quad S_i^2 = \frac{\sum_{j=1}^{n_i} (Y_{ij} - \bar{Y}_i)^2}{n_i - 1} \quad i = 1, 2 \quad S_p^2 = \frac{(n_1 - 1)S_1^2 + (n_2 - 1)S_2^2}{n_1 + n_2 - 2}$$

$$\bar{Y}_i \sim N\left(\mu_i, \frac{\sigma^2}{n_i}\right) \quad \frac{(n_i - 1)S_i^2}{\sigma^2} \sim \chi_{n_i - 1}^2$$

$$Z = \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sqrt{\sigma^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \sim N(0, 1) \quad W = \frac{(n_1 + n_2 - 2)S_p^2}{\sigma^2} \sim \chi_{n_1 + n_2 - 2}^2 \quad Z \perp W$$

$$\sqrt{\frac{W}{df_W}} = \sqrt{\frac{\frac{(n_1 + n_2 - 2)S_p^2}{\sigma^2}}{n_1 + n_2 - 2}} = \sqrt{\frac{S_p^2}{\sigma^2}} \quad \frac{Z}{\sqrt{\frac{W}{df_W}}} = \frac{\frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sqrt{\sigma^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}}}{\sqrt{\frac{S_p^2}{\sigma^2}}} = \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sqrt{S_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \sim t_{n_1 + n_2 - 2}$$

$$\Rightarrow P\left(-t_{\alpha/2, n_1 + n_2 - 2} \leq \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sqrt{S_p^2 \left(\frac{1}{n_1} + \frac{1}{n_2}\right)}} \leq t_{\alpha/2, n_1 + n_2 - 2}\right) = 1 - \alpha$$

Case 2 – Variances are Unequal - I

$$Y_{11}, \dots, Y_{1n_1} \sim N(\mu_1, \sigma_1^2) \quad Y_{21}, \dots, Y_{2n_2} \sim N(\mu_2, \sigma_2^2) \quad \sigma_1^2 \neq \sigma_2^2 \quad \{Y_1\} \perp \{Y_2\}$$

$$\frac{(n_i - 1)S_i^2}{\sigma_i^2} \sim \chi_{n_i - 1}^2 \Rightarrow E\left[\frac{(n_i - 1)S_i^2}{\sigma_i^2}\right] = n_i - 1, \quad V\left[\frac{(n_i - 1)S_i^2}{\sigma_i^2}\right] = 2(n_i - 1) \Rightarrow E(S_i^2) = \sigma_i^2, \quad V(S_i^2) = \frac{2\sigma_i^4}{n_i - 1}$$

$$Z = \frac{(\bar{Y}_1 - \bar{Y}_2) - (\mu_1 - \mu_2)}{\sqrt{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)}} \sim N(0, 1)$$

Problem : Replacing Denominator with estimated variances, consider :

$$\frac{W^*}{df_{W^*}} = \frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}\right)}{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)} \quad \text{which is NOT a chi - square divided by its degrees of freedom.}$$

$$\text{Aside: } W \sim \chi_\nu^2 \Rightarrow E(W) = \nu, \quad V(W) = 2\nu \Rightarrow E\left(\frac{W}{\nu}\right) = 1, \quad V\left(\frac{W}{\nu}\right) = \frac{2}{\nu}$$

$$E\left(\frac{W^*}{df_{W^*}}\right) = E\left(\frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}\right)}{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)}\right) = \frac{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)}{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)} = 1 \quad V\left(\frac{W^*}{df_{W^*}}\right) = \frac{1}{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)^2} \left\{ \frac{1}{n_1^2} \frac{2\sigma_1^4}{n_1 - 1} + \frac{1}{n_2^2} \frac{2\sigma_2^4}{n_2 - 1} \right\} = \frac{2}{\nu^*}$$

Case 2 – Variances are Unequal - II

$$V\left(\frac{W^*}{df_{W^*}}\right) = \frac{1}{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)^2} \left\{ \frac{1}{n_1^2} \frac{2\sigma_1^4}{n_1 - 1} + \frac{1}{n_2^2} \frac{2\sigma_2^4}{n_2 - 1} \right\} = \frac{2}{\nu^*} \Rightarrow$$

$$\frac{2}{\nu^*} = \frac{2\left(\frac{\sigma_1^4}{n_1^2(n_1 - 1)} + \frac{\sigma_2^4}{n_2^2(n_2 - 1)}\right)}{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)^2} \Rightarrow \nu^* = \frac{\left(\frac{\sigma_1^2}{n_1} + \frac{\sigma_2^2}{n_2}\right)^2}{\left(\frac{(\sigma_1^2/n_1)^2}{n_1 - 1} + \frac{(\sigma_2^2/n_2)^2}{n_2 - 1}\right)}$$

Replacing the unknown variances with their estimates:

$$\hat{\nu}^* = \frac{\left(\frac{S_1^2}{n_1} + \frac{S_2^2}{n_2}\right)^2}{\left(\frac{(S_1^2/n_1)^2}{n_1 - 1} + \frac{(S_2^2/n_2)^2}{n_2 - 1}\right)} = \frac{\left(\sum_{i=1}^2 g_i MS_i\right)^2}{\sum_{i=1}^2 \frac{(g_i MS_i)^2}{\nu_i}} \quad \text{where: } g_i = \frac{1}{n_i} \quad MS_i = S_i^2 \quad \nu_i = n_i - 1$$

So, we have the approximate degrees of freedom if our denominator were the square root of the ratio of a chi-square to its degrees of freedom

Example – Art Instruction Effect on Reading

- Experiment to Determine Effect of Art Instruction on a Reading Development Test
 - 52 Children Given Baseline Reading Test
 - 26 Received Art Instruction (Trt), 26 Did not (Control)
 - $Y = \text{Post-Test} - \text{Pre-Test Score}$

$$\bar{Y}_T = 7.77 \quad S_T^2 = 70.49 \quad n_T = 26$$

$$\bar{Y}_C = -1.58 \quad S_C^2 = 26.00 \quad n_C = 26$$

$$H_0: \sigma_T^2 = \sigma_C^2 \quad H_A: \sigma_T^2 \neq \sigma_C^2 \quad T.S.: F_{obs} = \frac{S_T^2}{S_C^2} = 2.71 \quad RR: \max(F_{obs}, 1/F_{obs}) \geq F_{.025, 25, 25} = 2.23$$

$$H_0: \mu_T = \mu_C \quad H_A: \mu_T \neq \mu_C \quad T.S.: t_{obs} = \frac{\bar{Y}_T - \bar{Y}_C}{\sqrt{\frac{S_T^2}{n_T} + \frac{S_C^2}{n_C}}} = \frac{7.77 - (-1.58)}{\sqrt{\frac{70.49}{26} + \frac{26.00}{26}}} = 4.85$$

$$\hat{\nu}^* = \frac{\left(\frac{70.49}{26} + \frac{26}{26}\right)^2}{\left(\frac{(70.49/26)^2}{25} + \frac{(26.00/26)^2}{25}\right)} = \frac{13.77}{0.33} = 41.23 \quad RR: |t_{obs}| \geq t_{.025, 41.23} = 2.020$$