Estimation of Static Discrete Choice Models Using Market Level Data

NBER Methods Lectures

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July 2012

Data Structures

- Market-level data
  - cross section/time series/panel of markets

- Consumer level data
  - cross section of consumers
  - sometimes: panel (i.e., repeated choices)
  - sometimes: second choice data

- Combination
  - sample of consumers plus market-level data
  - quantity/share by demographic groups
  - average demographics of purchasers of good j

Market-level Data

- We see product-level quantity/market shares by "market"
- Data include:
    - aggregate (market-level) quantity
    - prices/characteristics/advertising
    - definition of market size
    - distribution of demographics
        - sample of actual consumers
        - data to estimate a parametric distribution
- Advantages:
    - easier to get
    - sample selection less of an issue
- Disadvantages
    - estimation often harder and identification less clear

Consumer-level Data

- See match between consumers and their choices
- Data include:
    - consumer choices (including choice of outside good)
    - prices/characteristics/advertising of *all* options
    - consumer demographics
- Advantages:
    - impact of demographics
    - identification and estimation
    - dynamics (especially if we have panel)
- Disadvantages
    - harder/more costly to get
    - sample selection and reporting error

Review of the Model and Notation

Indirect utility function for the J inside goods

$$U(x_{jt}, \xi_{jt}, I_i - p_{jt}, D_{it}, \nu_{it}; \theta),$$

where

x_{jt} – observed product characteristics

ξ_{jt} – unobserved (by us) product characteristic

D_{it} – "observed" consumer demographics (e.g., income)

ν_{it} – unobserved consumer attributes

- ξ_{jt} will play an important role
    - realistic
    - will act as a "residual" (why don't predicted shares fit exactly – overfitting)
    - potentially implies endogeneity

Linear RC (Mixed) Logit Model

A common model is the linear Mixed Logit model

$$u_{ijt} = x_{jt}\beta_i + \alpha_i p_{jt} + \xi_{jt} + \varepsilon_{ijt}$$

where

$$\left(\begin{array}{c} \alpha_i \\ \beta_i \end{array} \right) = \left(\begin{array}{c} \alpha \\ \beta \end{array} \right) + \Pi D_i + \Sigma v_i$$

It will be coinvent to write

$$u_{ijt} \equiv \delta(x_{jt}, p_{jt}, \xi_{jt}; \theta_1) + \mu(x_{jt}, p_{jt}, D_i, v_i; \theta_2) + \varepsilon_{ijt}$$

where $\delta_{jt} = x_{jt}\beta + \alpha p_{jt} + \xi_{jt}$, and $\mu_{ijt} = (p_{jt} \ x_{jt})(\Pi D_i + \Sigma v_i)$

Linear RC (Mixed) Logit Model

A common model is the linear Mixed Logit model

$$u_{ijt} = x_{jt}\beta_i + \alpha_i p_{jt} + \xi_{jt} + \varepsilon_{ijt}$$

and

$$u_{ijt} \equiv \delta(x_{jt}, p_{jt}, \xi_{jt}; \theta_1) + \mu(x_{jt}, p_{jt}, D_i, \nu_i; \theta_2) + \varepsilon_{ijt}$$

Note:
(1) the mean utility will play a key role in what follows
(2) the interplay between μ_{ijt} and ε_{ijt}
(3) the "linear" and "non-linear" parameters
(4) definition of a market

Key Challenges for Estimation

- Recovering the non-linear parameters, which govern heterogeneity, without observing consumer level data
- The unobserved characteristic, ξ_{jt}
  - a main difference from early DC model (earlier models often had option specific constant in consumer level models)
  - generates a potential for correlation with price (or other x's)
  - when constructing a counterfactual we will have to deal with what happens to ξ_{jt}
- Consumer-level vs Market-level data
  - with consumer data, the first issue is less of a problem
  - the "endogeneity" problem can exist with both consumer and market level data: a point often missed

What would we do if we had micro data?

- Estimate in two steps.
- First step, estimate (δ, θ_2) say by MLE

$$\Pr(y_{it} = j | D_{it}, \mathbf{x}_t, \mathbf{p}_t, \boldsymbol{\xi}_t, \theta) = \Pr(y_{it} = j | D_{it}, \delta(\mathbf{x}_t, \mathbf{p}_t, \boldsymbol{\xi}_t, \theta_1), \mathbf{x}_t, \mathbf{p}_t, \theta$$

  e.g., assume ε_{ijt} is *iid* double exponential (Logit), and $\Sigma = 0$

$$= \frac{\exp\{\delta_{jt} + (p_{jt} \; x_{jt})\Pi D_i\}}{\sum_{k=0}^{J} \exp\{\delta_{kt} + (p_{kt} \; x_{kt})\Pi D_i\}}$$

- Second step, recover θ_1

$$\widehat{\delta}_{jt} = x_{jt}\beta + \alpha p_{jt} + \xi_{jt}$$

  ξ_{jt} is the residual. If it is correlated with price (or x's) need IVs (or an assumption about the panel structure)

Intuition from estimation with consumer data

- Estimation in 2 steps: first recover δ and θ_2 (parameters of heterogeneity) and then recover θ_1

- Different variation identfying the different parameters

    - θ_2 is identified from variation in demographics holding the level (i.e., δ) constant

        - If $\Sigma \neq 0$ then it is identified from within market share variation in choice probabilities

    - θ_1 is identified from cross market variation (and appropriate exclusion restrictions)

- With market-level data will in some sense try to follow a similar logic

    - however, we do not have within market variation to identify θ_2
    - will rely on cross market variation (in choice sets and demographics) for both steps
    - a key issue is that $\tilde{\xi}_{jt}$ is not held constant

In principle, we could consider estimating θ by min the distance between observed and predicted shares:

$$\min_{\theta} \| S_t - s_j(\mathbf{x}_t, \mathbf{p}_t, \theta) \|$$

- Issues:
  - computation (all parameters enter non-linearly)
  - more importantly,
    - prices might be correlated with the ξ_{jt} ("structural" error)
    - standard IV methods do not work

Inversion

- Instead, follow estimation method proposed by Berry (1994) and BLP (1995)
- Key insight:
  - with ξ_{jt} predicted shares can equal observed shares

  $$\sigma_j(\delta_t, \mathbf{x}_t, \mathbf{p}_t; \theta_2) = \int \mathbf{1}\left[u_{ijt} \geq u_{ikt} \quad \forall k \neq j\right] dF\left(\varepsilon_{it}, D_{it}, \nu_{it}\right) = S_{jt}$$

  - under weak conditions this mapping can be inverted

  $$\delta_t = \sigma^{-1}(\mathbf{S}_t, \mathbf{x}_t, \mathbf{p}_t; \theta_2)$$

  - the mean utility is linear in ξ_{jt}; thus, we can form linear moment conditions
  - estimate parameters via GMM

Important (and often missed) point

- IVs play dual role (recall 2 steps with consumer level data)
  - generate moment conditions to identify θ_2
  - deal with the correlation of prices and error
- Even if prices are exogenous still need IVs
- This last point is often missed
- Why different than consumer-level data?
  - with aggregate data we only know the mean choice probability, i.e., the market share
  - with consumer level data we know more moments of the distribution of choice probabilities (holding ξ_{jt} constant) : these moments help identify the heterogeneity parameters

- I will now go over the steps of the estimation
- For now I assume that we have valid IVs
  - later we will discuss where these come from
- I will follow the original BLP algorithm
  - I will discuss recently proposed alternatives later
  - I will also discuss results on the performance of the algorithm

The BLP Estimation Algorithm

1. Compute predicted shares: given δ_t and θ_2 compute $\sigma_j(\delta_t, \mathbf{x}_t, \mathbf{p}_t; \theta_2)$

2. Inversion: given θ_2 search for δ_t that equates $\sigma_j(\delta_t, \mathbf{x}_t, \mathbf{p}_t; \theta_2)$ and the observed market shares

   - the search for δ_t will call the function computed in (1)

3. Use the computed δ_t to compute ζ_{jt} and form the GMM objective function (as a function of θ)

4. Search for the value of θ that minimizes the objective function

Example: Estimation of the Logit Model

- Data: aggregate quantity, price characteristics. Market share $s_{jt} = q_{jt}/M_t$
  - Note: need for data on market size
- Computing market share

$$s_{jt} = \frac{\exp\{\delta_{jt}\}}{\sum_{k=0}^{J} \exp\{\delta_{kt}\}}$$

- Inversion

$$\ln(s_{jt}) - \ln(s_{0t}) = \delta_{jt} - \delta_{0t} = x_{jt}\beta + \alpha p_{jt} + \xi_{jt}$$

- Estimate using linear methods (e.g., 2SLS) with $\ln(s_{jt}) - \ln(s_{0t})$ as the "dependent variable".

Step 1: Compute the market shares predicted by the model

- Given δ_t and θ_2 (and the data) compute

$$\sigma_j(\delta_t, \mathbf{x}_t, \mathbf{p}_t; \theta_2) = \int \mathbf{1}\left[u_{ijt} \geq u_{ikt} \quad \forall k \neq j\right] \, dF\left(\varepsilon_{it}, D_{it}, \nu_{it}\right)$$

- For some models this can be done analytically (e.g., Logit, Nested Logit and a few others)
- Generally the integral is computed numerically
- A common way to do this is via simulation

$$\widetilde{\sigma}_j(\delta_t, x_t, p_t, F_{ns}; \theta_2) = \frac{1}{ns} \sum_{i=1}^{ns} \frac{\exp\{\delta_{jt} + (p_{jt}\ x_{jt})(\Pi D_i + \Sigma \nu_i)\}}{1 + \sum_{k=1}^{J} \exp\{\delta_{kt} + (p_{kt}\ x_{kt})(\Pi D_i + \Sigma}$$

  where ν_i and D_i, $i = 1, ..., ns$ are draws from $F_\nu^*(\nu)$ and $F_D^*(D)$,

- Note:
  - the ε's are integrated analytically
  - other simulators (importance sampling, Halton seq)
  - integral can be approximated in other ways (e.g., quadrature)

Step 2: Invert the shares to get mean utilities

- Given θ_2, for each market compute mean utility, δ_t, that equates the market shares computed in Step 1 to observed shares by solving

$$\widetilde{\sigma}(\delta_t, \mathbf{x}_t, \mathbf{p}_t, F_{ns}; \theta_2) = S_t$$

- For some model (e.g., Logit and Nested Logit) this inversion can be computed analytically.

- Generally solved using a contraction mapping for each market

$$\delta_t^{h+1} = \delta_t^h + \ln(S_t) - \ln(\widetilde{\sigma}(\delta_t^h, \mathbf{x}_t, \mathbf{p}_t, F_{ns}; \theta_2) \quad h = 0, ..., H,$$

where H is the smallest integer such that $\left\| \delta_t^H - \delta_t^{H-1} \right\| < \rho$

- δ_t^H is the approximation to δ_t

- Choosing a high tolerance level, ρ, is crucial (at least 10^{-12})

Step 3: Compute the GMM objective

- Once the inversion has been computed the error term is defined as

$$\xi_{jt}(\theta) = \widetilde{\sigma}^{-1}(\mathbf{S}_t, \mathbf{x}_t, \mathbf{p}_t; \theta_2) - x_{jt}\beta - \alpha p_{jt}$$

- Note: θ_1 enters this term, and the GMM objective, in a linear fashion, while θ_2 enters non-linearly.

- This error is interacted with the IV to form

$$\xi(\theta)' ZWZ' \xi(\theta)$$

where W is the GMM weight matrix

Step 4: Search for the parameters that maximize the objective

- In general, the search is non-linear
- It can be simplified in two ways.
    - "concentrate out" the linear parameters and limit search to θ_2
    - use the Implicit Function Theorem to compute the analytic gradient and use it to aid the search
- Still highly non-linear so much care should be taken:
    - start search from different starting points
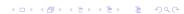   - use different optimizers

Identification

- Ideal experiment: randomly vary prices, characteristics and availability of products, and see where consumers switch (i.e., shares of which products respond)
- In practice we will use IVs that try to mimic this ideal experiment
- Next lecture we will see examples
- Is there "enough" variation to identify substitution?
- Solutions:
    - supply information (BLP)
    - many markets (Nevo)
    - add micro information (Petrin, MicroBLP)
- For further discussion and proofs (in NP case) see Haile and Berry

The Limit Distribution for the Parameter Estimates

- Can be obtained in a similar way to any GMM estimator
- With one cross section of observations is

$$J^{-1}(\Gamma'\Gamma)^{-1}\Gamma' V_0 \Gamma(\Gamma'\Gamma)^{-1}$$

- where
  - Γ – derivative of the expectation of the moments wrt parameters
  - V_0 – variance-covariance of those moments evaluated
- V_0 has (at least) two orthogonal sources of randomness
  - randomness generated by random draws on ξ
  - variance generated by simulation draws.
  - in samples based on a small set of consumers: randomness in sample
- Berry Linton and Pakes, (2004 RESTUD): last 2 components likely to be very large if market shares are small.
- A separate issue: limit in J or in T
  - in large part depends on the data

Challenges

- Knittel and Metaxoglou found that different optimizers give different results and are sensitive to starting values

- Some have used these results to argue against the use of the model

- Note, that its unlikely that a researcher will mimic the KM exercise
  - start from one starting point and not check others
  - some of the algorithms they use are not very good and rarely used

- It ends up that much (but not all!) of their results go away if
  - use tight tolerance in inversion (10^{-12})
  - proper code
  - reasonable optimizers

- This is an important warning about the challenges of NL estimation

MPEC

- Judd and Su, and Dube, Fox and Su, advocate the use of MPEC algorithm instead of the Nested fixed point
- The basic idea: maximize the GMM objective function subject to the "equilibrium" constraints (i.e., that the predicted shares equal the observed shares)
- Avoids the need to perform the inversion at each and every iteration of the search
    - performing the inversion for values of the parameters far from the truth can be quite costly
- The problem to solve can be quite large, but efficient optimzers (e.g., Knitro) can solve it effectively.
- DFS report significant speed improvements

MPEC (cont)

- Formally

$$\min_{\theta, \xi} \quad \xi' \mathbf{Z} \mathbf{W} \mathbf{Z}' \xi$$

$$\text{subject to} \quad \widetilde{\sigma}(\xi; \mathbf{x}, \mathbf{p}, F_{ns}, \theta) = S$$

- Note
    - the min is over both θ and ξ: a much higher dimension search
    - ξ is a vector of parameters, and unlike before it is not a function of θ
    - avoid the need for an inversion: equilibrium constraint only holds at the optimum
    - in principle, should yield the same solution as the nested fixed point

- Many bells and whistles that I will skip

ABLP

- Lee (2011) builds on some of the ideas proposed in dynamic choice to propose what he calls Approximate BLP
- The basic idea
  - Start with a guess to $\boldsymbol{\xi}$, denoted $\boldsymbol{\xi}^0$, and use it to compute a first order Taylor approximation to $\sigma(\boldsymbol{\xi}_t, \mathbf{x}_t, \mathbf{p}_t; \theta)$ given by

$$\ln s(\boldsymbol{\xi}_t; \theta) \approx \ln s^A(\boldsymbol{\xi}_t; \theta) \equiv \ln s(\boldsymbol{\xi}_t^0; \theta) + \frac{\partial \ln s(\boldsymbol{\xi}_t^0; \theta)}{\ln \boldsymbol{\xi}_t'}(\boldsymbol{\xi}_t - \boldsymbol{\xi}_t^0)$$

  - From $\ln S_t = \ln s^A(\boldsymbol{\xi}_t; \theta)$ we get

$$\boldsymbol{\xi}_t = \Phi_t(\boldsymbol{\xi}_t^0, \theta) \equiv \boldsymbol{\xi}_t^0 + \left[\frac{\partial \ln s(\boldsymbol{\xi}_t^0; \theta)}{\ln \boldsymbol{\xi}_t'} \right]^{-1} (\ln S_t - \ln s(\boldsymbol{\xi}_t^0; \theta))$$

  - Use this approximation for estimation

$$\min_{\theta} \Phi(\boldsymbol{\xi}^0, \theta)' \mathbf{ZWZ}' \Phi(\boldsymbol{\xi}^0, \theta)$$

ABLP (cont)

- Nest this idea into K-step procedure
  - Step 1: Obtain new GMM estimate

  $$\theta^K = \arg\min_{\theta} \Phi(\boldsymbol{\xi}^{K-1}, \theta)' \mathbf{ZWZ}' \Phi(\boldsymbol{\xi}^{K-1}, \theta)$$

  - Step 2: Update $\boldsymbol{\xi}$

  $$\boldsymbol{\xi}^K = \Phi(\boldsymbol{\xi}_t^{K-1}, \theta^K)$$

  - Repeat until convergence
- Like MPEC avoids inversion at each stage, but has low dimensional search
- Lee reports significant improvements over MPEC
- Disclaimer: still a WP and has not been significantly tested

Comparing the methods

Patel (2012, chapter 2 NWU thesis) compared the 3 methods

Table 12. Conditional on convergence, the average time in seconds until convergence from θ_C

	(a) MPEC			(b) NFP			(c) ABLP		
	Alternatives			Alternatives			Alternatives		
Markets	15	25	50	15	25	50	15	25	50
50	2.9	4.7	8.5	2.91	4.71	13.05	4.2	4.9	8.2
100	4.4	6.0	40.3	4.54	8.38	20.12	5.0	6.8	11.6
250	7.0	11.1	38.7	9.49	14.93	50.00	8.1	11.6	25.5
1000	46.9	127.0	128.3	30.67	59.09	184.31	23.0	35.1	81.7
2500	64.6	209.7	954.4	70.65	133.12	512.34	46.2	74.8	196.5

Comparing the methods

Table 17. Conditional on convergence, the average time in seconds until convergence from θ_F

| | (a) MPEC | | | (b) NFP | | | (c) ABLP | | |
| | Alternatives | | | Alternatives | | | Alternatives | | |
Markets	15	25	50	15	25	50	15	25	50
50	12.8	11.9	17.4	3.90	5.94	16.12	5.1	5.6	9.5
100	34.8	11.2	39.7	5.77	10.25	26.04	5.9	7.7	14.0
250	13.9	24.2	320.2	12.26	19.81	69.28	9.8	13.5	30.1
1000	-	128.8	-	38.84	80.68	272.79	30.3	41.1	96.2
2500	764.3	831.5	-	95.51	171.98	662.68	60.9	89.0	235.7

Word of caution: MPEC results can probably be **significantly** improved with better implementation
This is just an example of what one might expect if asking a (good) RA to program these methods