## Chapter 2 Multiple Regression <br> (Part 2)

## 1 Analysis of Variance in multiple linear regression

Recall the model again

$$
Y_{i}=\underbrace{\beta_{0}+\beta_{1} X_{i 1}+\ldots+\beta_{p} X_{i p}}_{\text {predictable }}+\underbrace{\varepsilon_{i}}_{\text {unpredictable }}, i=1, \ldots, n
$$

For the fitted model $\hat{Y}_{i}=b_{0}+b_{1} X_{i 1}+\ldots+b_{p} X_{i p}$,

$$
Y_{i}=\hat{Y}_{i}+e_{i} \quad i=1, \ldots, n
$$



|  | deviation of | deviation of <br> $\hat{Y}_{i}=b_{0}+b_{1} X_{i 1}+\ldots+b_{p} X_{i p}$ | deviation of <br> $e_{i}=Y_{i}-\hat{Y}_{i}$ |
| :---: | :---: | :---: | :---: |
| obs | $Y_{i}$ | $\hat{Y}_{1}-\bar{Y}$ | $e_{1}-\bar{e}=e_{1}$ |
| 1 | $Y_{1}-\bar{Y}$ | $\hat{Y}_{2}-\bar{Y}$ | $e_{2}-\bar{e}=e_{2}$ |
| 2 | $Y_{2}-\bar{Y}$ | $\vdots$ | $\vdots$ |
| $\vdots$ | $\vdots$ | $\hat{Y}_{n}-\bar{Y}$ | $e_{n}-\bar{e}=e_{n}$ |
| n | $Y_{n}-\bar{Y}$ | $\sum_{i=1}^{n}\left(\hat{Y}_{i}-\bar{Y}\right)^{2}$ | $\sum_{i=1}^{n} e_{i}^{2}$ |
| Sum of | $\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}$ | Sum of squares | Sum of squares |
| squares | Total Sum |  |  |
|  | of squares | due to regression | of error/residuals |
|  | (SST) | (SSR) | (SSE) |

We have

$$
\underbrace{\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}}_{\text {SST }}=\underbrace{\sum_{i=1}^{n}\left(\hat{Y}_{i}-\bar{Y}\right)^{2}}_{\mathrm{SSR}}+\underbrace{\sum_{i=1}^{n} e_{i}^{2}}_{\mathrm{SSE}}
$$

[Proof:

$$
\begin{aligned}
\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2} & =\sum_{i=1}^{n}\left(\hat{Y}_{i}-\bar{Y}+Y_{i}-\hat{Y}_{i}\right)^{2} \\
& =\sum_{i=1}^{n}\left\{\left(\hat{Y}_{i}-\bar{Y}\right)^{2}+\left(Y_{i}-\hat{Y}_{i}\right)^{2}+2\left(\hat{Y}_{i}-\bar{Y}\right)\left(Y_{i}-\hat{Y}_{i}\right)\right\} \\
& =S S R+S S E+2 \sum_{i=1}^{n}\left(\hat{Y}_{i}-\bar{Y}\right)\left(Y_{i}-\hat{Y}_{i}\right) \\
& =S S R+S S E+2 \sum_{i=1}^{n}\left(\hat{Y}_{i}-\bar{Y}\right) e_{i} \\
& =S S R+S S E
\end{aligned}
$$

where $\sum_{i=1}^{n} \hat{Y}_{i} e_{i}=0$ and $\sum_{i=1}^{n} e_{i}=0$ are used, which follow from the Normal equations. ]

$$
S S T=\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}=\mathbf{Y}^{\prime} \mathbf{Y}-\frac{\mathbf{1}}{\mathbf{n}} \mathbf{Y}^{\prime} \mathbf{J} \mathbf{Y}=\mathbf{Y}^{\prime}\left(\mathbf{I}-\frac{\mathbf{1}}{\mathbf{n}} \mathbf{J}\right) \mathbf{Y}
$$

Degree of freedom? $\mathbf{n - 1}$ (with n being the number of observations)
-

$$
S S E=\sum_{i=1}^{n} e_{i}^{2}=\mathbf{e}^{\prime} \mathbf{e}=(\mathbf{Y}-\mathbf{X b})^{\prime}(\mathbf{Y}-\mathbf{X b})=\mathbf{Y}^{\prime}(\mathbf{I}-\mathbf{H}) \mathbf{Y}
$$

Degree of freedom? n-p-1 (with p+1 being the number of coefficients)

- Let $\mathbf{H}=\mathbf{X}\left(\mathbf{X}^{\prime} \mathbf{X}\right)^{-\mathbf{1}} \mathbf{X}$ and and $\mathbf{J}=\mathbf{1 1}^{\prime} / \mathbf{n}$. Note that

$$
\hat{\mathbf{Y}}=\mathbf{H Y}
$$

and by the fact $\sum_{i=1}^{n} e_{i}=0$ (see the normal equations),

$$
\overline{\hat{Y}}=\bar{Y}=\mathbf{1}^{\prime} \mathbf{Y} / n
$$

Thus

$$
\begin{aligned}
S S R & =(\hat{\mathbf{Y}}-\bar{Y})^{\prime} *(\hat{\mathbf{Y}}-\bar{Y})=\mathbf{Y}^{\prime}(\mathbf{H}-\mathbf{J} / n)^{\prime}(\mathbf{H}-\mathbf{J} / n)^{\prime} \mathbf{Y} \\
& =\mathbf{Y}^{\prime}(\mathbf{H}-\mathbf{J} / n) \mathbf{Y} .
\end{aligned}
$$

Degree of freedom? $\mathbf{p}$ (the number of variables).
[Another Proof 1

$$
\hat{\mathbf{Y}}-\bar{Y}=\mathbf{H Y}-\mathbf{1}^{\prime} / \mathbf{n Y}=(\mathbf{H}-\mathbf{J} / \mathbf{n}) \mathbf{Y} .
$$

[^0]Write $\mathbf{X}=\left(\mathbf{1} \vdots \mathbf{X}_{\mathbf{1}}\right)$. Then

$$
H\left(\mathbf{1} \vdots \mathbf{X}_{\mathbf{1}}\right)=\mathbf{X}\left(\mathbf{X}^{\top} \mathbf{X}\right)^{-\mathbf{1}} \mathbf{X}^{\top} \mathbf{X}=\mathbf{X}=\left(\mathbf{1} \vdots \mathbf{X}_{\mathbf{1}}\right)^{\top}
$$

Thus

$$
H\left(\mathbf{1} \vdots \mathbf{X}_{\mathbf{1}}\right)=\mathbf{1}
$$

Similarly, $\mathbf{1}^{\prime} \mathbf{H}=\mathbf{1}^{\prime}$. Thus

$$
(\mathbf{H}-\mathbf{J} / n)^{\prime}(\mathbf{H}-\mathbf{J} / n)^{\prime}=\mathbf{H}-\mathbf{J} / n \mathbf{H}-\mathbf{H J} / n+\mathbf{J} / n=\mathbf{H}-\mathbf{J} / n
$$

]

- It follows that

$$
S S T=S S R+S S E
$$

We further define

$$
M S R=\frac{S S R}{p} \quad \text { called regression mean square }
$$

$M S E=\frac{S S E}{n-p-1} \quad$ called error mean square (or mean squared error)

## 2 ANOVA table

| Source of Variation | SS | df | MS | F-statistic |
| :---: | :---: | :---: | :---: | :---: |
| Regression | $S S R=\mathbf{Y}^{\prime}(\mathbf{H}-\mathbf{J} / n) \mathbf{Y}$ | $p$ | $M S R=\frac{S S R}{p}$ | MSR/MSE |
| Error | $S S E=\mathbf{Y}^{\prime}(\mathbf{I}-\mathbf{H}) \mathbf{Y}$ | $n-p-1$ | $M S E=\frac{S S E}{n-p-1}$ |  |
| Total | $S S T=\mathbf{Y}^{\prime}(\mathbf{I}-\mathbf{J} / n) \mathbf{Y}$ | $n-1$ |  |  |

## $3 \quad F$ test for regression relation

- $H_{0}: \beta_{1}=\beta_{2}=\ldots=\beta_{p}=0$ versus $H_{a}$ : not all $\beta_{k}(k=1, \ldots, p)$ equal zero
- Under $H_{0}$, the reduced model: $Y_{i}=\beta_{0}+\varepsilon_{i}$

$$
S S E(R)=S S T=\sum_{i=1}^{n}\left(Y_{i}-\bar{Y}\right)^{2}
$$

degrees of freedom $n-1$

- Full model: $Y_{i}=\beta_{0}+\beta_{1} X_{i 1}+\ldots+\beta_{p} X_{i p}+\varepsilon_{i}$

$$
S S E(F)=S S E=e^{\prime} e=(\mathbf{Y}-\mathbf{X} b)^{\prime}(\mathbf{Y}-\mathbf{X} b)
$$

degrees of freedom $n-p-1$

- $F$ test statistic (also called F-test for the model)

$$
F^{*}=\frac{(S S E(R)-S S E(F)) /(d f(R)-d f(F))}{S S E(F) / d f(F)}=\frac{S S R / p}{S S E /(n-p-1)}
$$

- If $F^{*} \leq F(1-\alpha ; p, n-p-1)$, conclude(accept) $H_{0}$ IF $F^{*}>F(1-\alpha ; p, n-p-1)$, conclude $H_{a}\left(\right.$ reject $\left.H_{0}\right)$


## $4 \quad R^{2}$ and the adjusted $R^{2}$

- $S S R=S S T-S S E$ is the part of variation explained by regression model
- Thus, define coefficient of multiple determination

$$
R^{2}=\frac{S S R}{S S T}=1-\frac{S S E}{S S T}
$$

which is the proportion of variation in the response that can be explained by the regression model (or that can be explained by the predictors $X_{1}, \ldots, X_{p}$ linearly)

- $0 \leq R^{2} \leq 1$
- with more predictor variables, SSE is smaller and $R^{2}$ is larger. To evaluate the contribution of the predictors fair, we define the adjusted $R^{2}$ :

$$
R_{a}^{2}=1-\frac{\frac{S S E}{n-p-1}}{\frac{S S T}{n-1}}=1-\left(\frac{n-1}{n-p-1}\right) \frac{S S E}{S S T}
$$

More discussion will be given later about $R_{a}^{2}$.

- For two models with the same number of predictor variables, $R^{2}$ can be used to indicate which model is better.
- If model A include more predictor variables than model B , then the $R^{2}$ of A must be equal or greater than that of model B. In that case, it is better to use the adjusted $R^{2}$.


## 5 Dwaine studios example

- $Y$-sales, $X_{1}$ - number of persons aged 16 or less, $X_{2^{-}}$income
- $n=21, p=3$
- $\mathrm{SST}=26,196.21, \mathrm{SSE}=2,180.93, \mathrm{SSR}=26,196.21-2,180.93=24,015.28$
- $F^{*}=\frac{24,015.28 / 2}{2,180.93 / 18}=99.1$

For $H_{0}: \beta_{1}=\beta_{2}=0$ with $\alpha=0.05, F(0.95 ; 2,18)=3.55$. because

$$
F^{*}>F(0.95 ; 2,18)
$$

we reject $H_{0}$
-

$$
R^{2}=\frac{24,015.28}{26.196 .21}=0.917, \quad R_{a}^{2}=0.907
$$

Writing a fitted regression model

|  | Coefficients: |  |  |  |  |
| :--- | :---: | :---: | :---: | :---: | :---: |
| Estimate | Std. Error | t value | $\operatorname{Pr}(>\|t\|)$ |  |  |
| (Intercept) | -68.8571 | 60.0170 | -1.147 | 0.2663 |  |
| x1 | 1.4546 | 0.2118 | 6.868 | $2 \mathrm{e}-06$ | $* * *$ |
| x2 | 9.3655 | 4.0640 | 2.305 | 0.0333 | $*$ |
| Residual standard error: 11.01 on 18 degrees of freedom |  |  |  |  |  |
| Multiple R-squared: 0.9167, Adjusted R-squared: 0.9075 |  |  |  |  |  |
| F-statistic: 99.1 on 2 and 18 DF, p-value: $1.921 \mathrm{e}-10$ |  |  |  |  |  |

The fitted model is

$$
\begin{gathered}
\hat{Y} \\
(\text { S.E. })
\end{gathered}=\underset{(60.02)}{-68.86}+\underset{(0.21)}{1.45 X_{1}}+\underset{(4.06)}{9.937 X_{2}}
$$

$R^{2}=0.9167, \quad R_{a}^{2}=0.9075, \quad$ F-statistic: 99.1 on 2 and 18 DF,


[^0]:    ${ }^{1}$ please ignore this proof

