# ECE 20875 <br> Python for Data Science <br> David Inouye and Qiang Qiu 

(Adapted from material developed by Profs. Milind Kulkarni, Stanley Chan, Chris Brinton, David Inouye)

## regression

## inference

- Inference is one of the basic problems that we want to solve in data science
- Given a set of data that we know some facts about, what new conclusions can we draw, and with what certainty?
- We will investigate several approaches to drawing conclusions from given sets of data

- Over the next few lectures: Making predictions about new data points given existing data using linear regression


## linear regression

- Basic modeling problem: I want to identify a relationship between ...
- explanatory variables (i.e., the "inputs", often referred to as the features of a data point), and
- a target variable (i.e., some "output" quantity that we want to estimate)
- Can we learn what this relationship is?
- If we have a model for this relationship, we
 can use it to predict the target variable for new data points


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## linear regression

- Can we learn the model from the data?
- Note that the model does not match the data exactly!
- A model is (at best) a simplification of the realworld relationship
- What makes a good model?
- Minimizes observed error: How far the model deviates from the observed data

- Maximizes generalizability: How well the model is expected to hold up to unseen data


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## simple linear regression model

- The simple linear regression model has a single explanatory variable:

$$
y_{n}=a x_{n}+b+\epsilon_{n}, \quad n=1, \ldots, N
$$

- $y_{n}$ is the measured value of the target variable for the $n$th data point
- $a x_{n}+b$ is the estimated value of the target, based on the explanatory $x_{n}$
- Each $y_{n}$ is associated with a model prediction
 component $a x_{n}+b$ plus some error term $\epsilon_{n}$
- How do we minimize this error?


## minimizing error

- The mean squared error (MSE) for simple linear regression is

$$
E(a, b)=\frac{1}{N} \sum_{n=1}^{N}\left(y_{n}-\left(a x_{n}+b\right)\right)^{2}
$$

- Common error metric: We looked at already when we studied the choice of histogram bin widths
- We want to minimize $E$, denoted: $\min E(a, b)$

$$
a, b
$$

- With two model parameters $a$ and $b$, this is
 reasonably easy to carry out by hand
- The square makes it easy to take the derivative


## minimizing error: derivation

- Set the derivatives with respect to $a$ and $b$ to zero:

$$
\begin{aligned}
& \frac{d E}{d a}=\frac{1}{N} \sum_{n=1}^{N}-2 x_{n}\left(y_{n}-\left(a x_{n}+b\right)\right)=0 \\
& \frac{d E}{d b}=\frac{1}{N} \sum_{n=1}^{N}-2\left(y_{n}-\left(a x_{n}+b\right)\right)=0
\end{aligned}
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\end{aligned}
$$

- Focusing first on the second equation, we have:

$$
\begin{aligned}
& \frac{-\sum_{n=1}^{N} y_{n}}{N}+a \frac{\sum_{n=1}^{N} x_{n}}{N}+b \frac{\sum_{n=1}^{N} 1}{N}=0, \text { or } \\
& b=\frac{\sum_{n=1}^{N} y_{n}}{N}-a \frac{\sum_{n=1}^{N} x_{n}}{N}=\bar{y}-a \bar{x}
\end{aligned}
$$

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$$

$$
a \frac{\sum_{n=1}^{N} x_{n}^{2}}{N}=\frac{\sum_{n=1}^{N} x_{n} y_{n}}{N}-b \frac{\sum_{n=1}^{N} x_{n}}{N}=\frac{\sum_{n=1}^{N} x_{n} y_{n}}{N}-b \bar{x}
$$

## minimizing error: derivation

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\frac{d E}{d b}=\frac{1}{N} \sum_{n=1}^{N}-2\left(y_{n}-\left(a x_{n}+b\right)\right)=0
\end{gathered}
$$

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$\frac{-\sum_{n=1}^{N} y_{n}}{N}+a \frac{\sum_{n=1}^{N} x_{n}}{N}+b \frac{\sum_{n=1}^{N} 1}{N}=0$, or
$b=\frac{\sum_{n=1}^{N} y_{n}}{N}-a \frac{\sum_{n=1}^{N} x_{n}}{N}=\bar{y}-a \bar{x}$
- As for the first equation,

$$
\begin{aligned}
& \frac{-\sum_{n=1}^{N} x_{n} y_{n}}{N}+a \frac{\sum_{n=1}^{N} x_{n}^{2}}{N}+b \frac{\sum_{n=1}^{N} x_{n}}{N}=0, \text { so } \\
& a \frac{\sum_{n=1}^{N} x_{n}^{2}}{N}=\frac{\sum_{n=1}^{N} x_{n} y_{n}}{N}-b \frac{\sum_{n=1}^{N} x_{n}}{N}=\frac{\sum_{n=1}^{N} x_{n} y_{n}}{N}-b \bar{x}
\end{aligned}
$$

- Substituting our expression for $b$, we have:

$$
\begin{aligned}
& a \frac{\sum_{n=1}^{N} x_{n}^{2}}{N}=\frac{\sum_{n=1}^{N} x_{n} y_{n}}{N}-(\bar{y}-a \bar{x}) \bar{x}, \text { or } \\
& a\left(\frac{\sum_{n=1}^{N} x_{n}^{2}}{N}-\bar{x}^{2}\right)=\frac{\sum_{n=1}^{N} x_{n} y_{n}}{N}-\bar{y} \bar{x}
\end{aligned}
$$

## minimizing error: formulas

- Isolating $a$ on the left hand side and simplifying, we get:
$a=\frac{\sum_{n=1}^{N} x_{n} y_{n}-N \bar{y} \bar{x}}{\sum_{n=1}^{N} x_{n}^{2}-N \bar{x}^{2}}$
- Here, $\bar{x}$ and $\bar{y}$ are the averages of the $x_{n}$ and $y_{n}$, respectively
- We can then use $a$ to solve for $b$ according to:
$b=\bar{y}-a \bar{x}$
- And then our linear regression predictor for a new
 datapoint $i$ is
$y_{i}=a x_{i}+b$


## minimizing error: formulas

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$$
y_{i}=a x_{i}+b
$$

- What do we do if there is more than one explanatory variable?
- To generalize to this case, it is more convenient to work with matrix equations


## matrix algebra review

- Let's say $\mathbf{x}=\left(\begin{array}{llll}x_{1} & x_{2} & \cdots & x_{n}\end{array}\right)^{T}$ and $\mathbf{y}=\left(\begin{array}{llll}y_{1} & y_{2} & \cdots & y_{n}\end{array}\right)^{T}$ are both $n$-dimensional vectors. Then

$$
\mathbf{x}^{T} \mathbf{y}=x_{1} y_{1}+x_{2} y_{2}+\cdots+x_{n} y_{n}
$$

is the inner product or dot product of $\mathbf{x}$ and $\mathbf{y}$, which is the multiplication of a $1 \times n$ and $n \times 1$ vector and results in a scalar.

- For example, suppose $\mathbf{x}=\left(\begin{array}{lll}3 & 4 & 5\end{array}\right)^{T}, \mathbf{y}=\left(\begin{array}{lll}1 & 0 & 2\end{array}\right)^{T}$. Then:

$$
\mathbf{x}^{T} \mathbf{y}=\left(\begin{array}{lll}
3 & 4 & 5
\end{array}\right)\left(\begin{array}{l}
1 \\
0 \\
2
\end{array}\right)=3 \times 1+4 \times 0+5 \times 2=13
$$

- The L2-norm of a vector $\mathbf{x}=\left(x_{1} x_{2} \cdots x_{n}\right)^{T}$ is a generalization of the Pythagorean theorem for finding the "length":

$$
\|\mathbf{x}\|_{2}=\sqrt{x_{1}^{2}+x_{2}^{2}+\cdots+x_{n}^{2}}
$$

## matrix algebra review

- More generally, define two $m \times n$ matrices:

$$
\mathbf{X}=\left[\begin{array}{cccc}
x_{11} & x_{12} & \cdots & x_{1 n} \\
x_{21} & x_{22} & \cdots & x_{2 n} \\
\vdots & \vdots & \ddots & \vdots \\
x_{m 1} & x_{m 2} & \cdots & x_{m n}
\end{array}\right], \quad \mathbf{Y}=\left[\begin{array}{cccc}
y_{11} & y_{12} & \cdots & y_{1 n} \\
y_{21} & y_{22} & \cdots & y_{2 n} \\
\vdots & \vdots & \ddots & \vdots \\
y_{m 1} & y_{m 2} & \cdots & y_{m n}
\end{array}\right]
$$

Then the matrix multiplication of $\mathbf{X}^{T}$ and $\mathbf{Y}$, which results in an $n \times n$ matrix, is:

$$
\mathbf{X}^{T} \mathbf{Y}=\left[\begin{array}{lll}
\mathbf{x}_{1} & \mathbf{x}_{2} & \cdots
\end{array} \mathbf{x}_{n}\right]^{T}\left[\begin{array}{lll}
\mathbf{y}_{1} & \mathbf{y}_{2} & \cdots
\end{array} \mathbf{y}_{n}\right]=\left[\begin{array}{c}
\mathbf{x}_{1}^{T} \\
\mathbf{x}_{2}^{T} \\
\vdots \\
\mathbf{x}_{n}^{T}
\end{array}\right]\left[\begin{array}{llll}
\mathbf{y}_{1} & \mathbf{y}_{2} & \cdots & \mathbf{y}_{n}
\end{array}\right]=\left[\begin{array}{cccc}
\mathbf{x}_{1}^{T} \mathbf{y}_{1} & \mathbf{x}_{1}^{T} \mathbf{y}_{2} & \cdots & \mathbf{x}_{1}^{T} \mathbf{y}_{n} \\
\mathbf{x}_{2}^{T} \mathbf{y}_{1} & \mathbf{x}_{2}^{T} \mathbf{y}_{2} & \cdots & \mathbf{x}_{2}^{T} \mathbf{y}_{n} \\
\vdots & \vdots & \ddots & \vdots \\
\mathbf{x}_{n}^{T} \mathbf{y}_{1} & \mathbf{x}_{n}^{T} \mathbf{y}_{2} & \cdots & \mathbf{x}_{n}^{T} \mathbf{y}_{n}
\end{array}\right]
$$

- For example, with $\mathbf{A}$ and $\mathbf{B}$ defined below, we get:

$$
\mathbf{A}=\left[\begin{array}{ccc}
-1 & 0 & 1 \\
0 & 2 & 3
\end{array}\right], \mathbf{B}=\left[\begin{array}{lll}
1 & 2 & 3 \\
3 & 0 & 1
\end{array}\right] \quad \rightarrow \quad \mathbf{A}^{T} \mathbf{B}=\left[\begin{array}{cc}
-1 & 0 \\
0 & 2 \\
1 & 3
\end{array}\right]\left[\begin{array}{lll}
1 & 2 & 3 \\
3 & 0 & 1
\end{array}\right]=\left[\begin{array}{ccc}
-1 & -2 & -3 \\
6 & 0 & 2 \\
10 & 2 & 6
\end{array}\right]
$$

## matrix algebra review

- If $\mathbf{X}$ has dimension $a \times b$, and $\mathbf{Y}$ has dimension $c \times d$, then the matrix product $\mathbf{X Y}$ is only possible if $b=c$ (i.e., the inner dimensions match), which will have dimension $a \times d$ (outer dimensions)
- If $\mathbf{X}$ is a square matrix (i.e., has dimension $n \times n$ ), then its inverse is $\mathbf{X}^{-1}$ (if it exists), and:

$$
\mathbf{X}^{-1} \mathbf{X}=\mathbf{X X}^{-1}=\mathbf{I}, \text { where } \mathbf{I}=\left[\begin{array}{cccc}
1 & 0 & \cdots & 0 \\
0 & 1 & \cdots & 0 \\
\vdots & \vdots & \ddots & \vdots \\
0 & 0 & \cdots & 1
\end{array}\right]
$$

is the $n \times n$ identity matrix

- For example, with $\mathbf{A}$ and $\mathbf{B}$ defined as below, we can verify $\mathbf{B}=\mathbf{A}^{-1}$, since $\mathbf{A B}=\mathbf{I}$ :

$$
\mathbf{A}=\left[\begin{array}{ccc}
3 & 0 & 2 \\
2 & 0 & -2 \\
0 & 1 & 1
\end{array}\right], \quad \mathbf{B}=\left[\begin{array}{ccc}
0.2 & 0.2 & 0 \\
-0.2 & 0.3 & 1 \\
0.2 & -0.3 & 0
\end{array}\right], \quad \mathbf{A} \mathbf{B}=\left[\begin{array}{lll}
1 & 0 & 0 \\
0 & 1 & 0 \\
0 & 0 & 1
\end{array}\right]
$$

## numpy

- But how do we perform matrix manipulations, like taking inverses, on large matrices in general?
- In Python, we can use the numpy library to do matrix operations

import numpy as np
np.array(A) //Convert list to numpy array
np.matmul(A,B) //Matrix multiplication (or A@B)
np.linalg.inv(A) //Matrix inverse
A.sum(axis=0) //Sum over rows of matrix
- See https://scipy-lectures.org/intro/numpy/operations.html for more examples, as well as the notebook


## matrix form of linear regression equations

- Now, back to regression
- For simple linear regression, if we define
$\mathbf{X}=\left[\begin{array}{cc}x_{1} & 1 \\ x_{2} & 1 \\ \vdots & \vdots \\ x_{N} & 1\end{array}\right] \quad \beta=\left[\begin{array}{l}a \\ b\end{array}\right] \quad \mathbf{y}=\left[\begin{array}{c}y_{1} \\ y_{2} \\ \vdots \\ y_{N}\end{array}\right]$
then we can write the equations for all data points compactly using the following matrix equation:

$$
\mathbf{y}=\mathbf{X} \beta+\epsilon
$$

- The multivariable linear regression model with $M$ explanatory variables is

$$
y_{n}=a_{1} x_{n, 1}+a_{2} x_{n, 2}+\cdots+a_{M} x_{n, M}+b+\epsilon_{n}, \quad n=1, \ldots, N
$$

- In this case, we define

$$
\mathbf{X}=\left[\begin{array}{ccccc}
x_{1,1} & x_{1,2} & \cdots & x_{1, M} & 1 \\
x_{2,1} & x_{2,2} & \cdots & x_{2, M} & 1 \\
\vdots & \vdots & \ddots & \vdots & \vdots \\
x_{N, 1} & x_{N, 2} & \cdots & x_{N, M} & 1
\end{array}\right] \quad \beta=\left[\begin{array}{c}
a_{1} \\
a_{2} \\
\vdots \\
a_{M} \\
b
\end{array}\right] \quad \mathbf{y}=\left[\begin{array}{c}
y_{1} \\
y_{2} \\
\vdots \\
y_{N}
\end{array}\right]
$$

where $\mathbf{X}$ is the feature matrix. Then, as before, we can write

$$
\mathbf{y}=\mathbf{X} \beta+\epsilon
$$

## least squares equations

- With this matrix notation, we can write our original optimization for minimizing MSE as:

$$
\min _{\beta} \frac{1}{N} \sum_{n=1}^{N}\left(y_{n}-\mathbf{x}_{n}^{T} \beta\right)^{2}
$$

- Or, equivalently, this can be written using the vector norm:

$$
\min _{\beta} \frac{1}{N}\|\mathbf{y}-\mathbf{X} \beta\|_{2}^{2}
$$

- Similar to 1D case, we can take the gradient (multidimensional derivative) and set to 0 (i.e., the vector of zeros) to find minimum:
$\nabla\left((1 / N)\|\mathbf{y}-\mathbf{X} \beta\|_{2}^{2}\right)=(2 / N) \mathbf{X}^{T} \mathbf{X} \beta-(2 / N) \mathbf{X}^{T} \mathbf{y}=\mathbf{0}$
- This yields the least squares equations for solving for $\beta$ :
$\mathbf{X}^{T} \mathbf{X} \beta=\mathbf{X}^{T} \mathbf{y}$


## solving for $\beta$

- If $\mathbf{X}^{T} \mathbf{X}$ is invertible, we can take a matrix inverse to solve for the model parameters $\beta$ :

$$
\beta=\left(\mathbf{X}^{T} \mathbf{X}\right)^{-1} \mathbf{X}^{T} \mathbf{y}
$$

- But $\mathbf{X}^{T} \mathbf{X}$ is not always invertible
- The inverse exists if and only if the columns of $\mathbf{X}$ are linearly independent of one another
- This means that we cannot have the case where one column can be written as a linear combination of the others
- What does it mean when $\mathbf{X}^{T} \mathbf{X}$ is not invertible?

- Infinitely many possible solutions
- We typically choose the one where $\|\beta\|$ is smallest. Why?


## example

Suppose we collect five data points consisting of two features $x_{1}, x_{2}$ and a target variable $y$ in the form $\left(x_{1}, x_{2}, y\right)$ : $(1,2,10),(-3,6,0),(0,0,3),(1,-1,4)$, $(5,-2,20)$. We want to fit a linear regression model to this dataset.

What are the least squares equations?
What is the resulting model?
What would be the prediction for a new datapoint with $x_{1}=-1, x_{2}=1$ ?

## solution: least squares equations

The model we want to fit is $\hat{y}=a_{1} x_{1}+a_{2} x_{2}+b$, where $\beta=\left(a_{1} a_{2} b\right)^{T}$ is the parameter vector. The feature matrix $\mathbf{X}$, target vector $\mathbf{y}$, and least squares equations are:

$$
\begin{gathered}
\mathbf{X}=\left[\begin{array}{ccc}
1 & 2 & 1 \\
-3 & 6 & 1 \\
0 & 0 & 1 \\
1 & -1 & 1 \\
5 & -2 & 1
\end{array}\right], \quad \mathbf{y}=\left(\begin{array}{c}
10 \\
0 \\
3 \\
4 \\
20
\end{array}\right], \\
{\left[\begin{array}{ccccc}
1 & -3 & 0 & 1 & 5 \\
2 & 6 & 0 & -1 & -2 \\
1 & 1 & 1 & 1 & 1
\end{array}\right]\left[\begin{array}{ccc}
1 & 2 & 1 \\
-3 & 6 & 1 \\
0 & 0 & 1 \\
1 & -1 & 1 \\
5 & -2 & 1
\end{array}\right] \beta=\left[\begin{array}{ccccc}
1 & -3 & 0 & 1 & 5 \\
2 & 6 & 0 & -1 & -2 \\
1 & 1 & 1 & 1 & 1
\end{array}\right]\left(\begin{array}{c}
10 \\
0 \\
3 \\
4 \\
20
\end{array}\right)} \\
\mathbf{X}^{T} \mathbf{X} \beta=\mathbf{X}^{T} \mathbf{y}
\end{gathered}
$$

## solution: model and test prediction

Using the numpy commands for inverse, transpose, and multiplication, we compute the solution: $\beta=\left(\mathbf{X}^{T} \mathbf{X}\right)^{-1} \mathbf{X}^{T} \mathbf{y}$

$$
\beta=(4.2308,1.7538,2.2615)^{T}
$$

Which means that our model is

$$
\hat{y}=4.2308 x_{1}+1.7538 x_{2}+2.2615
$$

And the prediction for $x_{1}=-1, x_{2}=1$ is

$$
\hat{y}=4.2308 \cdot-1+1.7538 \cdot 1+2.2615=-0.2154
$$

## interpreting results

- How should we interpret the results of linear regression?
- Recall multi-feature model, e.g.,

$$
y_{n}=a_{1} x_{n, 1}+a_{2} x_{n, 2}+b
$$

- If one feature weight (e.g., $a_{1}$ ) is higher than another (e.g., $a_{2}$ ), this can indicate that this feature is more important than the other (contributes more to the value of $y$ )
- Need to be careful, though! If different features have different scales, then weights will naturally be different!


Here, $x_{1}$ has a range of 8 , while $x_{2}$ only has a range of 2

- Normalization is useful as it standardizes the feature ranges


## normalization for interpretation

- Problem: Suppose I fit a linear regression model and get

$$
\hat{y}=10 x_{1}+100 x_{2}+5
$$

- Does this mean that $x_{2}$ has a bigger impact on $y$ than $x_{1}$ ?
- Not necessarily, because we have said nothing about the ranges of $x_{1}$ and $x_{2}$ that resulted in $a_{1}=10$ and $a_{2}=100$.
- One solution: Normalize the data before doing linear regression so that coefficients are comparable over a consistent range.


## standard normalization

- For every feature column, do the following to make them all have a mean of 0 and standard deviation of 1 :

1. Center values: Subtract the column average from each feature sample

- Useful to eliminate any bias contained in the features

2. Scale values: Divide each feature sample by the column standard deviation

- Re-scales features so that each is expressed in new units: standard deviations from the mean (similar to how we calculate $z$-scores)
- Mathematically, we are defining the following operation for each feature column $\mathbf{x}_{m}$ :
$\tilde{\mathbf{x}}_{m}=\frac{\mathbf{x}_{m}-\bar{x}_{m}}{s_{m}}$, where $\bar{x}_{m}$ and $s_{m}$ are the sample mean and standard deviation of feature $m$


## coefficient of determination

- How good is the fit of the regression to the dataset?
- To answer this, one possibility is using the MSE
- Another commonly used quantity is the coefficient of determination, called $r^{2}$

$$
r^{2}=1-\frac{\sum_{n=1}^{N}\left(y_{n}-\hat{y}_{n}\right)^{2}}{\sum_{n=1}^{N}\left(y_{n}-\bar{y}\right)^{2}}=1-\frac{M S E}{\sigma_{Y}^{2}}
$$

- $y_{n}$ : Measured value, $\hat{y}_{n}$ : Predicted value
- $\bar{y}$ : Mean measured value, $\sigma_{Y}^{2}$ : Variance of measured value
- $r^{2}$ gives the fraction of variance in the data that is explained by the model
- Typically between 0 (bad, no better than horizontal line) and 1 (perfect fit)
- Sometimes preferred to MSE in regression problems for this reason



## using your model after fitting

- After fitting a linear regression model, you can estimate (or predict) the target $y$ of new data points using your model
- New data point: $\left(x_{1}, x_{2}, \ldots\right)$
- Prediction: $\hat{y}=a_{1} x_{1}+a_{2} x_{2}+\cdots+b$
- How good is the prediction?
- Squared error between $\hat{y}$ and $y$ (once it is known)
- MSE or $r^{2}$ over a set of new data points
- When using the model, make sure to take into account any normalization that was used (i.e., normalize new data points before inputting them, "un-normalize" the $\hat{y}$ you get back)



## linear regression in python

- You can solve the least squares equations directly using numpy
- Given how common linear regression is, several variants are built in to the sklearn (scikit learn) library directly:

```
from sklearn import linear_model, from sklearn.metrics import
mean_squared_error, r2_score
regr = linear_model.LinearRegression(fit_intercept=True) # Define
linear regression object
regr.fit(X_train,y_train) # Fit model to training set
regr.coef_ # View coefficients (a_1,...,a_M) of trained model
regr.intercept_ # View intercept (b) of trained model
y_pred = regr.predict(X_test) # Apply model to test set
r2_score(y_true,y_pred) # r2 score between true and predicted
```


## more interpretation

- Is a feature significant?
- Just because a feature is used in a model doesn't mean it is important in predicting the value of the output
- But the model will try to account for the feature anyway!
- Can perform a hypothesis test (see previous lectures):
- Null hypothesis $H_{0}$ : Coefficient $a_{m}$ is 0 (feature has no predictivity, $y$ does not depend on $x_{m}$ )
- Alternative hypothesis $H_{1}$ : Coefficient $a_{m}$ is not 0 (feature has predictivity, $y$ does depend on $x_{m}$ )



## hypothesis test for regression

- Test statistic is always: (value - hypothesized value) / standard error

$$
\frac{\hat{a}_{m}-a_{m}}{S E_{a_{m}}} \Rightarrow \frac{\hat{a}_{m}}{S E_{a_{m}}}
$$



$$
S E_{a_{m}}=\frac{\sqrt{\frac{\sum_{n=1}^{N}\left(y_{n}-\hat{y}_{n}\right)^{2}}{N-2}}}{\sqrt{\sum_{n=1}^{N}\left(x_{n, m}-\bar{x}_{m}\right)^{2}}}
$$

- $y_{n}$ : Measured value, $x_{n, m}$ : Feature value
- $\hat{y}_{n}$ : Predicted value, $\bar{x}_{m}$ : Feature average
- For a $z$-test, find $p$-value of $S E_{a_{m}}$ against the $z$-distribution
- For a $t$-test, find $p$-value against a $t$-distribution with $N-k-1$ degrees of freedom, where $k$ is the number of features



## a linear model may be wrong

- In these graphs, all 4 datasets have the same...
- linear regression line
- coefficient of determination
- mean and variance of both $x$ and $y$
- Yet clearly, the relationship between x and $y$ is different in each case
- It is important to visualize the results, and possibly try non-linear models!






## what about non-linear?

- A common (and understandable) misconception is that linear regression can only find linear relationships
- The "linear" part refers to the parameter vector $\beta$, not the input features in $\mathbf{X}$
- We can readily take nonlinear functions of our features
- For example, suppose we want to fit a quadratic model:

$$
y_{n}=a_{1}\left(x_{n}\right)^{2}+a_{2} x_{n}+b
$$

- We create a "synthesized" feature matrix that has the quadratic form:

$$
\mathbf{X}=\left[\begin{array}{ccc}
\left(x_{1}\right)^{2} & x_{1} & 1 \\
\left(x_{2}\right)^{2} & x_{2} & 1 \\
\vdots & \vdots & \vdots \\
\left(x_{N}\right)^{2} & x_{N} & 1
\end{array}\right] \quad \beta=\left[\begin{array}{c}
a_{1} \\
a_{2} \\
b
\end{array}\right] \quad \mathbf{y}=\left[\begin{array}{c}
y_{1} \\
y_{2} \\
\vdots \\
y_{N}
\end{array}\right]
$$



## more and more complexity

- If we use a higher degree $d$ of polynomials, we can reduce MSE:



- But, is this a good thing to do?





## overfitting

- If our goal was just to minimize error on the existing dataset, we'd keep adding features (e.g., increasing the degree $d$ of a polynomial)
- But this sacrifices the generalizability of the model
- An overfitted model is one which contains too many parameters than can be justified by the data
- High $r^{2}$ and low MSE on training data, but low $r^{2}$ and high MSE on testing data
- We can contrast this with underfitting, where we don't have enough parameters to drive down MSE on either
 training or testing data


## regularization

- When we have a lot of features, we can use regularization, a class of techniques for mitigating overfitting by penalizing non-zero model coefficients
- The general expression we work with in regularization is:

```
minimize (model error) + }\lambda\mathrm{ (coefficient weights)
```

- $\lambda \geq 0$ is the regularization parameter
- Higher $\lambda$ : Minimizing model parameters becomes more important
- Lower $\lambda$ : Minimizing model error becomes more important
- Several different regularization techniques: Lasso, Ridge, Elastic-Net, ...


## ridge regression

- In ridge regression, the regularization term is the sum of squares of the coefficients:

$$
\underset{\beta}{\operatorname{minimize}}\|\mathbf{X} \beta-\mathbf{y}\|_{2}^{2}+\lambda\|\beta\|_{2}^{2}
$$

- This makes it easy to solve in matrix form as:

$$
\beta^{\star}=\left(\mathbf{X}^{T} \mathbf{X}+\lambda \mathbf{I}\right)^{-1} \mathbf{X}^{T} \mathbf{y}
$$

- In Python (where $\alpha$ is the regularization parameter):


```
from sklearn import linear_model
reg = linear_model.Ridge(alpha=0.1, fit_intercept=True)
```


## regularization can alleviate overfitting

- Polynomial of degree $d=10$, with different amounts of regularization:





 has a "smoothing" effect on the model


## evaluating predictive performance

- Descriptive and diagnostic analysis (classical statistics, data mining)
- Focus: Understand and interpret statistical relationships in observed dataset
- Evaluation: e.g., MSE or $r^{2}$ on training data (data used to fit the model)
- Predictive and prescriptive analysis (machine learning)
- Focus: Predict target value for new or future unseen data
- Evaluation: e.g., MSE or $r^{2}$ on test data (data not used to fit the model)


## why evaluate on test data?

- Analogy to class
- Training data is like homeworks, sample problems and sample exams
- Testing data is like the real exam
- If we train and evaluate on the same data, the model may not generalize well
- Reasons for computing performance on test data (the standard ML approach):
- Model evaluation: Quantify the model's predictive performance if deployed
- e.g., describing the model and its business implications to the CEO
- Model selection: Select which model should be deployed

- e.g., which polynomial degree or regularization value should be used?


## choosing model based on test MSE

- We can use MSE on a held-out test set to determine the best model:


MSE $=0.04$
Test MSE $=0.04$


MSE $=0.64$
Test MSE $=0.44$


MSE $=0.01$
Test MSE $=0.14$


- Blue points: Training set
- Orange points: Held-out test set


## choosing model based on test MSE

- We can use MSE on a held-out test set to determine the best model:
 the lowest test MSE
- This is often achieved when there is a small difference between training and test MSE


## simulating testing data

- Ultimately, we'd like to actually test the model in the real world (e.g., predict tomorrow's temperature)
- However, this is usually quite costly, time consuming, or downright impossible, so we have to simulate it
- To do this, we can split our dataset into:
- Training data: A subset we use to train/fit the model

- Testing data: A subset we used to report the generalized performance
- Common splits: 90/10 (i.e., $90 \%$ training and $10 \%$ test) and $80 / 20$
- Note: It is important that the algorithm never sees the testing data (just like it is important that students don't see the real midterm)


## cross validation

- $k$-fold cross validation (often abbreviated CV) repeats the train/ test split idea $k$ times, across different folds of the data
- The data is divided into $k$ parts
- In each fold, one part is used as the testing set, and the other $k-1$ are used as the training set
- Thus, there are $k$ models fit throughout this process, and we can average testing performance (and sometimes the coefficients)
- How many folds should be used?
- 3-fold, 5-fold and 10-fold are common



## Average

- Leave-one-out CV: $k$ is the number of datapoints, i.e., one is held out in each fold (computationally expensive)


## cross validation for model selection

- How do we determine the right value of $\lambda$ ?
- Test a wide range of $\lambda$ typically log scale, e.g., $0.01, \ldots, 0.1, \ldots, 1, \ldots, 10, \ldots, 100$
- Use multiple CV iterations, one for each value of $\lambda$ :

- Choose $\lambda^{\star}$ whose CV performance is the best
- For final model, train model with all data using $\lambda^{\star}$

Train all folds
with $\lambda=1$


Train all folds with $\lambda=10$



## (very small) cv example

Suppose we collect three data points with a single feature $x$ and target variable $y$. In the form ( $x, y$ ), they are, approximately: $(2.18,2.26),(0.13$, -14.57), (2.75, 16.74).

Find the linear regression model $\hat{y}=a x+b$ and corresponding regularization parameter $\lambda$ which has minimum cross validation error.

Use the Ridge model, $k=3$ folds, and test $\lambda=0,0.1,1$. Note that the coefficient $b$ should NOT be regularized.

## solution

- We need to solve the least squares equations for three values of lambda, and three folds each (i.e., 9 cases total). Here is the math for $\lambda=0,0.1$ and the second fold:

```
fold=2, lambda=0.0
X:
[[2.17997451 1. ]
    [2.74831239 1. ] ]
X.T @ X:
[[12.30550986 4.9282869 ]
    [ 4.9282869 2. ]]
X.T @ X + lambda*I:
[[12.30550986 4.9282869 ]
    [ 4.9282869 2. ]]
(X.T @ X + lambda*I)^(-1):
[[ [ 6.19179817 -15.25747891]
    [-15.25747891 38.09661673]]
(X.T @ X + lambda*I)^(-1)@ X^T:
[[-1.75951672 1.75951672]
    [ 4.8357016 -3.8357016 ]]
(X.T @ X + lambda*I)^(-1)@ X^T @ Y:
[ 25.47215001 -53.26685674]
x ~ [2.18, 0.13, 2.75]
y ~ [2.26, -14.57, 16.74]

Only coefficient is changed by \(\lambda\), intercept is not regularized

Notice how different the inverse is just from a small \(\lambda\)
```

fold=2, lambda=0.1

```
fold=2, lambda=0.1
X:
```

X:

```
```

[[2.17997451 1. ]

```
[[2.17997451 1. ]
    [2.74831239 1. ] ]
    [2.74831239 1. ] ]
X.T @ X:
X.T @ X:
[[12.30550986 4.9282869 ]
[[12.30550986 4.9282869 ]
    [ 4.9282869 2. ]]
    [ 4.9282869 2. ]]
X.T @ X + lambda*I:
X.T @ X + lambda*I:
[[12.40550986 4.9282869 ]
[[12.40550986 4.9282869 ]
    [ 4.9282869 2. ]]
    [ 4.9282869 2. ]]
(X.T @ X + lambda*I)^(-1):
(X.T @ X + lambda*I)^(-1):
[[ 3.82403369 -9.42296757]
[[ 3.82403369 -9.42296757]
    [-9.42296757 23.71954383]]
    [-9.42296757 23.71954383]]
(X.T @ X + lambda*I)^(-1)@ X^T:
(X.T @ X + lambda*I)^(-1)@ X^T:
[[-1.0866716 1.0866716]
[[-1.0866716 1.0866716]
    [ 3.1777147 -2.1777147]]
    [ 3.1777147 -2.1777147]]
(X.T @ X + lambda*I)^(-1)@ X^T @ Y:
(X.T @ X + lambda*I)^(-1)@ X^T @ Y:
[ 15.73151403 -29.26453239]
```

[ 15.73151403 -29.26453239]

```
fold=2, lambda=0.0

X:
[ [2.17997451 1. ]
[2.74831239 1. ]]
X.T @ X:
\(\left[\begin{array}{ll}{[12.30550986} & 4.9282869]\end{array}\right]\)
[ 4.9282869 2. ]]
X.T @ \(\mathrm{X}+\mathrm{lambda*I:}\)
\(\left[\begin{array}{lll}{[12.30550986} & 4.9282869\end{array}\right]\)
[ 4.9282869 2. ]]
(X.T @ \(\mathrm{X}+\mathrm{lambda*}\) ) ^( -1 ):
[ \(\left[\begin{array}{lll}{[6.19179817-15.25747891]}\end{array}\right]\)
\(\left[\begin{array}{lll}-15.25747891 & 38.09661673]\end{array}\right]\) (X.T @ \(\mathrm{X}+\mathrm{lambda*I)}\) ^(-1)@ \(\mathrm{X}^{\wedge} \mathrm{T}:\) \(\left[\begin{array}{ll}{[-1.75951672} & 1.75951672\end{array}\right]\)
[ 4.8357016 -3.8357016 ] ]
(X.T @ X + lambda*I)^(-1) @ X^T @ y:
[ 25.47215001 -53.26685674]
fold=2, lambda=0.1
X:
[ [2.17997451 1. ]
[2.74831239 1. ]]
X.T @ X:
[ [12.30550986 4.9282869 ]
[ 4.92828692.
X.T @ \(\mathrm{X}+\mathrm{lambda*I:}\)
[ [12.40550986 4.9282869 ]
[ 4.9282869 2. ]]
(X.T @ X + lambda*I)^(-1):
[ [ 3.82403369 -9.42296757]
[-9.42296757 23.71954383]]
(X.T @ X + lambda*I)^(-1)@ \(\mathrm{X}^{\wedge} \mathrm{T}\) :
\(\left[\begin{array}{ll}{[-1.0866716} & 1.0866716\end{array}\right]\)
[ 3.1777147-2.1777147]]
(X.T @ X + lambda*I)^(-1)@ X^T @ \(\mathrm{y}:\)
[ 15.73151403 -29.26453239]
\(\lambda^{\star}=0.10\) has best average test MSE
```

