Introducing Linear Regression: An Example Using Basketball Statistics

Tom Arnold and Jonathan Godbey¹

ABSTRACT

The intuition behind linear regression can be difficult for students to grasp particularly without a readily accessible context. This paper uses basketball statistics to demonstrate the purpose of linear regression and to explain how to interpret its results. In particular, the student will quickly grasp the meaning of explanatory variables, r-squared, and the statistical significance of estimates of regression coefficients. Even if the student is not a sports fan the examples are easily understood and familiar. The student can easily replicate the procedures in this paper to reinforce learning.

Introduction

When calculators were introduced into the classroom, a number of tedious calculations could suddenly be performed very quickly. However, students' comprehension of mathematics did not actually improve and the calculator in many ways masked deficiencies because a keystroke sequence could substitute for comprehension. Regression analysis has some of the same characteristics because econometric software has improved to the point that a regression is a simple one line command that results in copious amounts of output.

We believe the reason for the difficulty in understanding/interpreting regression results is not the product of students being unable to perform the regression using matrix algebra, but the product of a lack of intuition that belies the one line of code. In other words, it is equally important to understand why the regression is being performed, what the regression process does to the data, and how to interpret the regression results.

By using the statistics from a basketball team, students become enabled to comprehend a model for predicting the number of points a given player should be able to score based on certain factors. Regression is then introduced as a means to calibrate and test the model.

The paper begins with a breakdown of what a regression "does" based on a very small set of data in which hand calculators can perform all of the calculations. Next, the statistics from a basketball team are presented so that a model can be generated to predict how many points should be scored in a game by an individual given a set of factors. A regression is performed to calibrate and test the model. A second regression based on the capital asset pricing model (Sharpe, 1964) is then performed on actual financial data. The paper concludes at this point.

What Actually Happens in a Regression?

Before introducing the basketball data, start with something even smaller. In Table 1, we have three columns of data: A, B, and C with the averages of each column calculated by summing the data and dividing it by the number of observations (5 in this case).

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Observation:	Data A:	Data B:	Data C:
1	22	10	8
2	47	21	15
3	34	19	12
4	21	8	5
5	66	25	19
Mean:	38.00	16.60	11.80

Table 1: Data

Assume existing theory states that A is dependent on B and C in some manner, usually expressed as a model. In other words, there is some combination of data B and data C that generates data A. Further assume, that theory views the relationship as linear: $A = \beta 1^*B + \beta 2^*C$. With this information: data and a proposed model of how B and C generate A, discuss with the class criteria for selecting $\beta 1$ and $\beta 2$. After some discussion, if it has not already been suggested, try the criteria: Mean (A) less Mean (B) less Mean (C) is zero (i.e. the mean of the error in the model is zero). Be certain to make the point: without this condition, on average, the model will be in error. Next, suggest some combinations of $\beta 1$ and $\beta 2$: (1.2000, 1.5322), (1.5783, 1.0000), and (2.5000, -0.2966). Each combination meets the criteria:

$0 = 38.00 - \{(1.2000)*16.60 + (1.5322)*11.80\}$	(1)
$0 = 38.00 - \{(1.5783)*16.60 + (1.0000)*11.80\}$	(2)
$0 = 38.00 - \{(2.5000) \times 16.60 + (-0.2966) \times 11.80\}$	(3)

However, of the three combinations, which one is the best?

Discussion usually leads to an answer of using the combination of β_1 and β_2 with the least amount of error. The problem with such a conclusion is that it depends on what is meant by error. On average, all three combinations have equivalent error, but the error is different for each combination (see Table 2).

Observation:	$\beta 1 = 1.2000$ $\beta 2 = 1.5322$	$\beta 1 = 1.5783$ $\beta 2 = 1.0000$	$\beta 1 = 2.5000$ $\beta 2 = -0.2966$	
1	-2.258	-1 783	-0.627	
2	-1.183	-1.144	-1.051	
3	-7.186	-7.988	-9.941	
4	3.739	3.374	2.483	
5	6.888	7.543	9.135	
Mean of Error:	0.00	0.00	0.00	

Table 2: Model Error

To say one version of error is better than another version of error is fairly difficult from observing the error. Also, unless a systematic rule is in place, there can be no consistent means of determining which version of the error is best. After some more discussion (note: the discussion is important to get the student to understand the issue and possibly resolve the issue), suggest looking at the square of the error and take the mean of the squared error (i.e. the mean squared error).

	$\beta 1 = 1.2000$	$\beta 1 = 1.5783$	$\beta 1 = 2.5000$
Observation:	$\beta 2 = 1.5322$	$\beta 2 = 1.0000$	$\beta 2 = -0.2966$
1	5.097	3.197	0.393
2	1.399	1.3089	1.105
3	51.644	63.803	98.820
4	13.980	11.381	6.165
5	47.447	56.889	83.456
Mean of Squared Error:	23.914	27.312	37.988

Table 3: Model Squared Error

When viewing the mean squared error, $\beta 1 = 1.2000$ and $\beta 2 = 1.5322$ appear to be the best performing values of the three combinations. However, suppose $\beta 1$ is set to 1.0040 and $\beta 2$ is set to 1.8400 (see Table 4).

Observation:	Error:		Squared Error:					
1	-2.760		7.618					
2	-1.684		2.836					
3	-7.156		51.208					
4	3.768		14.198					
5	5.940		35.284					
Mean of Error:	-0.378	Mean of Squared Error:	22.229					

Table 4: Model with $\beta 1 = 1.0040$ and $\beta 2 = 1.8400$

Notice, the mean squared error is reduced versus the best case from Table 3, however, the mean of the error is not zero. The question becomes: is it possible to use these two values for $\beta 1$ and $\beta 2$ to get a lower mean squared error and have the mean of the error be zero as well? After some discussion (note: again, the discussion needs to happen if the student is going internalize the intuition of the analysis), suggest using an intercept term α of -0.3784. Now the relationship becomes: $A = \alpha + \beta 1^*B + \beta 2^*C$. By allowing an intercept term, the mean squared error is further reduced and the mean of the model now equals zero. (4)

 $0 = 38.00 - \{-0.3784 + (1.0040)*16.60 + (1.8400)*11.80\}$

At this point, the intuition of what occurs in a regression calculation is complete. Regression is a means of finding the coefficient combination (that can include an intercept) that will minimize the squared error between the dependent variable (A) and a set of independent variables (B and C). By visualizing that different coefficient combinations generate different mean squared errors, the students can see that an optimal solution exists and that this is ultimately what the computer code for the regression accomplishes. Now the question becomes: what are examples of data A, data B, and data C?

A Simple Example

Consider the following situation. Georgia State University's basketball team has 12 players. Each player averages a different number of points per game (PPG). Table 5 shows all the information we have about the players: PPG, average minutes played per game (MPG), average rebounds per game (RPG), and jersey number (NUM).

Player	PPG	MPG	RBG	NUM
Dukes	12.8	32.1	4.7	2
Goldston	10.8	27.1	1.5	11
Mendez	8.8	27.8	2.9	21
Chase	5.1	25.6	5.8	15
Hansbro	4.9	14.8	2.8	23
Curry	6.9	19.1	1.6	12
Hampton	4.4	22.1	3.6	1
Krubally	2.9	10.9	2.9	24
Fields	2.3	11.5	1.1	4
Lott	2.2	10.9	1.2	30
Rimmer	2.5	13.3	2.7	33
Echols	7.4	21.6	5.6	0

Table 5: Player Information

If we did not know PPG, would it be possible to calculate the exact PPG of each player? If it is not possible to calculate the exact PPG, is an estimate possible? How good is that estimate? In other words does MPG, RPG or NUM explain part or all of PPG? Perhaps we only need to know one of the three variables to

estimate PPG. Maybe one is enough to get an approximate estimate but knowing one or both of the others will give us a more precise estimate. Linear regression provides the means to answer these questions.

Note that we only have 12 players. It makes sense that using more observations would make us more confident in our estimates. If we were truly trying to answer the question of what explains PPG we would want data from many other teams. However, for purposes of trying to "see" what is happening with the numbers it is better to keep the number of observations small.

Most basketball players would agree that playing more minutes means you will score more points. At least we are reasonably certain that playing more minutes will not mean that you score fewer points. Figure 1 shows the relationship between points scored and minutes played. Clearly, playing more minutes results in scoring more points. Mathematically, we want the following relationship.

 $PPG = \alpha + \beta * MPG$

Figure 1



If we know α and β , then all we need to know is MPG and we can easily calculate PPG. For example, assume α is 1 and β is 0.5. If a player plays 20 minutes then he must have scored 1 + 0.5*20 = 11 points. Unfortunately, we do not know the true α or β . We need to estimate them. Figure 1 shows that the relationship is not exact. The points do not lie on a straight line. There is no α and β that will allow us to determine the exact PPG for every player. There will be some error. The best we can do is find an α and β combination that will come as close as possible to estimating PPG using MPG. Following the intuition from the last section, a linear regression solves for the best combination of α and β (i.e., it minimizes the mean squared error) that allows MPG (the independent variable) to estimate PPG (the dependent variable). Our new equation (with the error term added) is:

$$PPG = \alpha + \beta * MPG + \varepsilon \tag{6}$$

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(5)

Fortunately, Excel does the calculations quickly and easily.

Regression in Excel

Rather than having an equation like equation (2), Excel uses the general form: $Y = \alpha + \beta * X + \varepsilon$. To run the regression, enter data as shown in Figure 2.

(7)

Figure	2
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2	Dukes	12.8	32.1	4.7	2									
3	Goldston	10.8	27.1	1.5	11									
4	Mendez	8.8	27.8	2.9	21									
5	Chase	5.1	25.6	5.8	15									
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11	Lott	2.2	10.9	1.2	30									
12	Rimmer	2.5	13.3	2.7	33									
13	Echols	7.4	21.6	5.6	0									
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Go to the Data tab, click Data Analysis, then click Regression. Now highlight cells B2 to B13 for the Y data, cells C2 to C13 for the X data, and choose a cell for output. Click OK and Excel will produce the results shown in Figure 3. (To make Figure 3 easier to read the initial output was re-formatted.) Note: Many times the Data Analysis tab is not available in an initial setup for Excel. However, it can be added by going to the File tab in Excel 2010, click Options, and click Analysis ToolPak. In the previous version of Excel, click the clover-leaf-like icon, click Excel options at the bottom of the menu, click Add-ins, and click Analysis ToolPak.

Interpreting Regression Results

Our question remains, does MPG explain PPG? If it does then β from equation (6) has to be different from zero. If β were zero then any number of MPG would result in the same estimate for PPG. In other words MPG would not explain PPG. Mathematically, α could take any value. However, we know in reality it must be zero. If a player averages zero MPG he must score zero PPG. Figure 3 gives the estimates for α and β .

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6	Adjusted R Square	0.774			0.000								
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10	ANOVA												
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12	Regression	1.000	106.820	106.820	38.764	0.00	00						
13	Residual	10.000	27.557	2.756									
14	Total	11.000	134.377										
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16		Coefficients	Standard Error	t Stat	P-value	Lower 95%	Upper 95%	Lower 95.0%	Upper 95.0%				
17	Intercept	-2.370	1.415	-1.675	0.125	-5.52	0.782	-5.522	0.782				
18	X Variable 1	0.420	0.067	6.226	0.000	0.2	0.570	0.270	0.570				
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Figure 3

The coefficient for "Intercept" is α and the coefficient for "X Variable 1" is β . These are estimates, not the true values of α and β . We now have the following.

PPG = -2.370 + 0.420 * MPG

The estimate for α was -2.370 when we know the true value to be zero. How likely is it that the estimate would be found to be zero? The p-value of 0.125 answers that question. This p-value means that there is a 12.5% chance that the estimated value of α is zero and equal to (in this case) the true value of zero. This also means there is an 87.5% chance that the estimated α is different from zero. Because we know α should be zero, this result does not seem to make sense. There are three explanations for this seemingly incorrect result. First, we have a small sample size. Perhaps more data would result in an estimate closer to zero. When adding two more teams (Drexel University and James Madison University), α becomes -0.9950 which is closer to zero than the initial regression, but ironically has a lower p-value (6.27%) as well. Perhaps even more data is needed or one of the remaining two explanations may need to be considered in addition.

Second, other variables may matter in addition to MPG which are currently omitted from the regression. The effect of these "omitted variables" would be captured in α . Additional variables will be considered later in the paper (RPG and NUM), however, take this opportunity to discuss with students what other variables may matter. One example is a binary variable such as home game versus away game. Such a variable will not work with the type of data presented here, but a very rich discussion about the quality of data can be had as result. Discuss with the students what kind of data is preferable...cumulative data (like the data presented here) or individual game data, or an even finer increment of data. Note: many times economic and financial data come in varying frequencies making this discussion very relevant.

Third, estimates are generally considered to be different from zero only if the p-value is 0.10 or less. If the p-value is greater than 0.10 then the estimate is said to be insignificant. If the p-value is less than 0.10, then the estimate is said to be significant at the 10% level. If the p-value is less than 0.05, then the estimate is said to be significant at the 5% level and less than 0.01 means the estimate is significant at the 1% level. In general, estimates are grouped into one of the four categories above.

In the initial regression with only GSU players, α is insignificant. However, with additional data, α becomes significant at the 10% level. In this context, α has to be zero (again, you cannot score points if you do

not play) and being estimated as different from zero while being insignificant is "statistically" the equivalent of being zero. The initial regression demonstrates this case. However, with additional data, the estimate for α gets closer to zero, but then becomes significant. Although an apparently "confounding result", this is not an uncommon situation and as discussed earlier leads to the desire for more data, more variables, or both to address the confounding result.

It should be noted, that many times the true value of α is not known with any certainty within a particular model and analyzing the estimated α is simply a matter of determining if the estimated α is different from zero. In this case, α is known to be zero which allowed for a more in depth discussion of the regression model not being fully accurate.

The estimate for β was 0.420. If the true value of β is different from zero then MPG is useful in explaining PPG. Is 0.420 close to zero? Once again, the p-value of 0.000098 answers that question. (Note Figure 3 shows the rounded result 0.000). This means that there is a 0.0098% chance that the true value of β is zero and our estimate is different from zero as a result of chance. In other words, we are 99.9902% sure that the true β is not zero. MPG matters. MPG is significant at the 1% level.

MPG matters, but does it explain all of PPG? Would knowing RPG (average rebounds per game) or NUM (jersey number) help in the estimation of PPG? The r-square given in Figure 3 indicates the percentage of the variation of PPG that is explained by MPG. The result of 0.795 means that 79.5% of the variation in PPG is explained by MPG (Note: this is an upwardly biased estimate and more analysis presented later in this paper can be used to address this issue). If the result had been 1.000 then MPG would explain 100% of the variation in PPG. Because not all of the variation in PPG is explained by MPG, let us consider searching for more explanatory variables.

Multivariate Regressions

The significance of multiple variables can be checked simultaneously. Suppose we want to see if RPG and NUM in addition to MPG are important in explaining PPG because we believe more variables are necessary to produce a better estimation of PPG. We, then, run the following regression.

 $PPG = \alpha + \beta 1 * MPG + \beta 2 * RPG + \beta 3 * NUM + \varepsilon$

(9)

To run this regression, enter data as shown in Figure 2. Go to the Data tab, click Data Analysis, then click Regression. Now highlight cells B2 to B13 for the Y data, cells C2 to E13 for the X data, and choose a cell for output. Click OK and Excel will produce the results shown in Figure 4. (To make Figure 4 easier to read the initial output was re-formatted.)

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4	Multiple R	0.91254													
5	R Square	0.83273													
6	Adjusted R Square	0.77001													
7	Standard Error	1.67619													
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Figure 4

Excel separates output based on the columns of the input. "X Variable 1" corresponds to our β 1, "X Variable 2" to our β 2 and "X Variable 3" to our β 3. If RPG and NUM are important in explaining PPG then the estimates for β 2 and β 3 should be significantly different from zero.

First, consider the intuition. What do you expect? Most people guess that RPG might be important but that NUM would not be. Logically, one could assume that a player who rebounds well is a talented player and has the ability to score but the number on a player's jersey will not add any ability to score, rebound or to perform any task on the court. A counterargument could be easily made. Rebounding is a separate skill. Those who are good at it score fewer points or those who miss a lot of shots must rebound as a result. Possibly, rebounding is a defensive skill and such players focus on their comparative advantage in rebounding instead of shooting for the good of the team. Scoring specialists score, rebounding specialists rebound. Therefore, players with more rebounds will score less. The estimate for $\beta 2$ should then be negative and significantly different from zero. It may be that the skills are separate and have no impact on one another. The estimate should then be insignificant.

It could be possible that jersey numbers matter. While not the source of superior performance the number could be a signal of it. Teams may assign lower numbers to better scorers. In other sports, jersey numbers actually do correspond with specific types of positions. No matter if one believes this to be true or not, it is a testable condition within a regression.

It is important to discuss with the class that the above arguments could be totally false, completely true or partial explanations. Let the class determine what data should be important and provide an explanation why the data is important for explaining PPG (Figure 4 reveals the truth about such contentions).

The coefficient of MPG is 0.462 and its p-value is 0.001. MPG is significant at the 1% level. We are 99.9% certain that the true coefficient is not zero. All else constant, playing another minute will result in scoring 0.462 more points. RPG and NUM have coefficients of -0.489 and -0.016, respectively. The p-values

are 0.218 and 0.760. Neither one is significant at the 10% level and therefore both are considered insignificant. In other words, they do not matter.

The intercept (α) is -1.482, which is closer to zero than when estimated in the initial regression with only MPG as an independent variable. Unlike the initial regression, one cannot definitively state that α should be zero because RPG and NUM have also been included in the regression. However, RPG and NUM are considered insignificant while MPG is still significant. Consequently, because RPG and NUM are insignificant, it again makes sense for α to be zero. As stated before, α is closer to zero than it had been in the initial regression, but what is more striking is how much more insignificant α has become: p-value = 0.511 versus a p-value of 0.125 previously. Again, "statistically", α is effectively estimated to be zero.

The r-square is 0.833 which is higher than the r-square of the regression containing only MPG. R-square will never fall if additional explanatory variables are added. It may even rise if additional insignificant explanatory variables are added. This result demonstrates a weakness of r-square. The addition RPG and NUM which do not matter in the estimation of PPG should not cause r-square to rise. To address this issue the measure "adjusted r-square" was developed. The adjusted r-square lowers r-square as more explanatory variables are added to the equation. This way, the effects of adding insignificant explanatory variables becomes mitigated.

Regression with Financial Data

In the same way we used linear regression to identify variables that explain scoring in basketball, we may use it to identify variables that explain the returns on individual stocks. We begin with the theoretical Capital Asset Pricing Model (CAPM) and ExxonMobil (XOM). The CAPM implies that the returns on XOM can be explained by returns on the market. We will assume the market is the S&P 500. The equation is:

 $E[r_{XOM}] = r_f + \beta * [E[r_{S\&P500}] - r_f].$

(10)

Equation (10) states that the expected return on XOM is the risk-free rate (r_f) plus some risk premium. That risk premium is β times the expected return on the S&P 500 above the risk-free rate. To test this relationship, first, subtract the risk-free rate from both sides. The result is: $E[r_{XOM}] - r_f = \alpha + \beta * [E[r_{S&P500}] - r_f].$ (11)

Returns for XOM are calculated using monthly prices as given by finance.yahoo.com from January 2000 to December 2008. Returns for the S&P 500 over the same time period are calculated using data from finance.yahoo.com. The risk-free rate is assumed to be the 3-month T-bill as shown on www.federalreserve.gov. The risk-free rate is subtracted from both XOM and S&P500 returns. Now run the regression

 $R_{XOM} = \alpha + \beta * R_{S\&P500} + \varepsilon.$

(12)

 R_{XOM} are the returns on XOM above the risk-free rate and shown in Figure 5, column G. $R_{S\&P500}$ are the returns on the S&P500 above the risk-free rate and shown in Figure 5, column H. To run the regression highlight G4..G122 for Y and H4..H122 for X.

Figure 5)
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3	1/3/2000	33.33		1,394.46		0.0532										=
4	2/1/2000	30.46	-0.0861	1,366.42	-0.0201	0.0555	-0.1416	-0.0756								
5	3/1/2000	31.55	0.0358	1,498.58	0.0967	0.0569	-0.0211	0.0398								
6	4/3/2000	31.42	-0.0041	1,452.43	-0.0308	0.0566	-0.0607	-0.0874								
7	5/1/2000	33.88	0.0783	1,420.60	-0.0219	0.0579	0.0204	-0.0798								
8	6/1/2000	31.92	-0.0579	1,454.60	0.0239	0.0569	-0.1148	-0.0330								
9	7/3/2000	32.61	0.0216	1,430.83	-0.0163	0.0596	-0.0380	-0.0759								
10	8/1/2000	33.38	0.0236	1,517.68	0.0607	0.0609	-0.0373	-0.0002								
11	9/1/2000	36.44	0.0917	1,436.51	-0.0535	0.0600	0.0317	-0.1135								
12	10/2/2000	36.47	0.0008	1,429.40	-0.0049	0.0611	-0.0603	-0.0660								
13	11/1/2000	36.16	-0.0085	1,314.95	-0.0801	0.0617	-0.0702	-0.1418								
14	12/1/2000	35.72	-0.0122	1,320.28	0.0041	0.0577	-0.0699	-0.0536								
15	1/2/2001	34.58	-0.0319	1,366.01	0.0346	0.0515	-0.0834	-0.0169								
16	2/1/2001	33.48	-0.0318	1,239.94	-0.0923	0.0488	-0.0806	-0.1411								
17	3/1/2001	33.46	-0.0006	1,160.33	-0.0642	0.0442	-0.0448	-0.1084								
18	4/2/2001	36.59	0.0935	1,249.46	0.0768	0.0387	0.0548	0.0381								
19	5/1/2001	36.84	0.0068	1,255.82	0.0051	0.0362	-0.0294	-0.0311								
20	6/1/2001	36.26	-0.0157	1,224.38	-0.0250	0.0349	-0.0506	-0.0599								
21	7/2/2001	34.67	-0.0438	1,211.23	-0.0107	0.0351	-0.0789	-0.0458								
22	8/1/2001	33.52	-0.0332	1,133.58	-0.0641	0.0336	-0.0668	-0.0977								
23	9/4/2001	32.90	-0.0185	1,040.94	-0.0817	0.0264	-0.0449	-0.1081								
24	10/1/2001	32.94	0.0012	1,059.78	0.0181	0.0216	-0.0204	-0.0035								
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25	11/1/2001	31.41	-0.0464	1,139.45	0.0752	0.0187	-0.0651	0.0565								1
26	12/3/2001	33.01	0.0509	1,148.08	0.0076	0.0169	0.0340	-0.0093								
27	1/2/2002	32.80	-0.0064	1,130.20	-0.0156	0.0165	-0.0229	-0.0321								
28	2/1/2002	34.89	0.0637	1,106.73	-0.0208	0.0173	0.0464	-0.0381								
29	3/1/2002	37.03	0.0613	1,147.39	0.0367	0.0179	0.0434	0.0188								
30	4/1/2002	33.94	-0.0834	1,076.92	-0.0614	0.0172	-0.1006	-0.0786								
31	5/1/2002	33.93	-0.0003	1,067.14	-0.0091	0.0173	-0.0176	-0.0264								
32	6/3/2002	34.77	0.0248	989.82	-0.0725	0.0170	0.0078	-0.0895								
33	7/1/2002	31.24	-0.1015	911.62	-0.0790	0.0168	-0.1183	-0.0958								
34	8/1/2002	30.32	-0.0294	916.07	0.0049	0.0162	-0.0456	-0.0113								
35	9/3/2002	27.28	-0.1003	815.28	-0.1100	0.0163	-0.1166	-0.1263								
36	10/1/2002	28.79	0.0554	885.76	0.0864	0.0158	0.0396	0.0706								
37	11/1/2002	29.96	0.0406	936.31	0.0571	0.0123	0.0283	0.0448								
38	12/2/2002	30.08	0.0040	879.82	-0.0603	0.0119	-0.0079	-0.0722								
39	1/2/2003	29.40	-0.0226	855.70	-0.0274	0.0117	-0.0343	-0.0391								
40	2/3/2003	29.48	0.0027	841.15	-0.0170	0.0117	-0.0090	-0.0287								
41	3/3/2003	30.29	0.0275	848.18	0.0084	0.0113	0.0162	-0.0029								
42	4/1/2003	30.51	0.0073	916.92	0.0810	0.0113	-0.0040	0.0697								
43	5/1/2003	31.77	0.0413	963.59	0.0509	0.0107	0.0306	0.0402								
44	6/2/2003	31.34	-0.0135	974.50	0.0113	0.0092	-0.0227	0.0021								
45	7/1/2003	31.06	-0.0089	990.31	0.0162	0.0090	-0.0179	0.0072								
46	8/1/2003	33.13	0.0666	1,008.01	0.0179	0.0095	0.0571	0.0084								
47	9/2/2003	32.17	-0.0290	995.97	-0.0119	0.0094	-0.0384	-0.0213								
48	10/1/2003	32.15	-0.0006	1,050.71	0.0550	0.0092	-0.0098	0.0458								
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49	11/3/2003	32.04	-0.0034	1,058.20	0.0071	0.0093	-0.0127	-0.0022								
50	12/1/2003	36.29	0.1326	1,111.92	0.0508	0.0090	0.1236	0.0418								
51	1/2/2004	36.10	-0.0052	1,131.13	0.0173	0.0088	-0.0140	0.0085								
52	2/2/2004	37.55	0.0402	1,144.94	0.0122	0.0093	0.0309	0.0029								
53	3/1/2004	37.04	-0.0136	1,126.21	-0.0164	0.0094	-0.0230	-0.0258								
54	4/1/2004	37.89	0.0229	1,107.30	-0.0168	0.0094	0.0135	-0.0262								
55	5/3/2004	38.76	0.0230	1,120.68	0.0121	0.0102	0.0128	0.0019								
56	6/1/2004	39.80	0.0268	1,140.84	0.0180	0.0127	0.0141	0.0053								
57	7/1/2004	41.50	0.0427	1,101.72	-0.0343	0.0133	0.0294	-0.0476								
58	8/2/2004	41.57	0.0017	1,104.24	0.0023	0.0148	-0.0131	-0.0125								
59	9/1/2004	43.58	0.0484	1,114.58	0.0094	0.0165	0.0319	-0.0071								
60	10/1/2004	44.38	0.0184	1,130.20	0.0140	0.0176	0.0008	-0.0036								=
61	11/1/2004	46.46	0.0469	1,173.82	0.0386	0.0207	0.0262	0.0179								
62	12/1/2004	46.47	0.0002	1,211.92	0.0325	0.0219	-0.0217	0.0106								
63	1/3/2005	46.78	0.0067	1,181.27	-0.0253	0.0233	-0.0166	-0.0486								
64	2/1/2005	57.67	0.2328	1,203.60	0.0189	0.0254	0.2074	-0.0065								
65	3/1/2005	54.30	-0.0584	1,180.59	-0.0191	0.0274	-0.0858	-0.0465								
66	4/1/2005	51.95	-0.0433	1,156.85	-0.0201	0.0278	-0.0711	-0.0479								
67	5/2/2005	51.46	-0.0094	1,191.50	0.0300	0.0284	-0.0378	0.0016								
68	6/1/2005	52.62	0.0225	1,191.33	-0.0001	0.0297	-0.0072	-0.0298								
69	7/1/2005	53.79	0.0222	1,234.18	0.0360	0.0322	-0.0100	0.0038								
70	8/1/2005	55.12	0.0247	1,220.33	-0.0112	0.0344	-0.0097	-0.0456								
71	9/1/2005	58.46	0.0606	1,228.81	0.0069	0.0342	0.0264	-0.0273								
72	10/3/2005	51.66	-0.1163	1,207.01	-0.0177	0.0371	-0.1534	-0.0548								
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73	11/1/2005	53.67	0.0389	1,249.48	0.0352	0.0388	0.0001	-0.0036		-		_			-	1
74	12/1/2005	51.95	-0.0320	1,248.29	-0.0010	0.0389	-0.0709	-0.0399								
75	1/3/2006	58.03	0.1170	1,280.08	0.0255	0.0424	0.0746	-0.0169								
76	2/1/2006	55.20	-0.0488	1,280.66	0.0005	0.0443	-0.0931	-0.0438								
77	3/1/2006	56.58	0.0250	1,294.87	0.0111	0.0451	-0.0201	-0.0340								
78	4/3/2006	58.65	0.0366	1,310.61	0.0122	0.0460	-0.0094	-0.0338								
79	5/1/2006	56.91	-0.0297	1,270.09	-0.0309	0.0472	-0.0769	-0.0781								
80	6/1/2006	57.33	0.0074	1,270.20	0.0001	0.0479	-0.0405	-0.0478								
81	7/3/2006	63.30	0.1041	1,276.66	0.0051	0.0495	0.0546	-0.0444								
82	8/1/2006	63.52	0.0035	1,303.82	0.0213	0.0496	-0.0461	-0.0283								
83	9/1/2006	62.99	-0.0083	1,335.85	0.0246	0.0481	-0.0564	-0.0235								
84	10/2/2006	67.04	0.0643	1,377.94	0.0315	0.0492	0.0151	-0.0177								
85	11/1/2006	72.42	0.0803	1,400.63	0.0165	0.0494	0.0309	-0.0329								
86	12/1/2006	72.25	-0.0023	1,418.30	0.0126	0.0485	-0.0508	-0.0359								
87	1/3/2007	69.86	-0.0331	1,438.24	0.0141	0.0498	-0.0829	-0.0357								
88	2/1/2007	67.87	-0.0285	1,406.82	-0.0218	0.0503	-0.0788	-0.0721								
89	3/1/2007	71.44	0.0526	1,420.86	0.0100	0.0494	0.0032	-0.0394								≡
90	4/2/2007	75.16	0.0521	1,482.37	0.0433	0.0487	0.0034	-0.0054								
91	5/1/2007	79.09	0.0523	1,530.62	0.0325	0.0473	0.0050	-0.0148								
92	6/1/2007	79.76	0.0085	1,503.35	-0.0178	0.0461	-0.0376	-0.0639								
93	7/2/2007	80.95	0.0149	1,455.27	-0.0320	0.0482	-0.0333	-0.0802								
94	8/1/2007	81.85	0.0111	1,473.99	0.0129	0.0420	-0.0309	-0.0291								
95	9/4/2007	88.37	0.0797	1,526.75	0.0358	0.0389	0.0408	-0.0031								
96	10/1/2007	87.82	-0.0062	1,549.38	0.0148	0.0390	-0.0452	-0.0242								
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97	11/1/2007	85.45	-0.0270	1.481.14	-0.0440	0.0327	-0.0597	-0.0767			K	-				
98	12/3/2007	89.79	0.0508	1,468.36	-0.0086	0.0300	0.0208	-0.0386								
99	1/2/2008	82.14	-0.0852	1,378.55	-0.0612	0.0275	-0.1127	-0.0887								
100	2/1/2008	83.75	0.0196	1,330.63	-0.0348	0.0212	-0.0016	-0.0560								
101	3/3/2008	81.41	-0.0279	1,322.70	-0.0060	0.0126	-0.0405	-0.0186								
102	4/1/2008	89.59	0.1005	1,385.59	0.0475	0.0129	0.0876	0.0346								
103	5/1/2008	85.82	-0.0421	1,400.38	0.0107	0.0173	-0.0594	-0.0066								
104	6/2/2008	85.21	-0.0071	1,280.00	-0.0860	0.0186	-0.0257	-0.1046								
105	7/1/2008	77.76	-0.0874	1,267.38	-0.0099	0.0163	-0.1037	-0.0262								
106	8/1/2008	77.75	-0.0001	1,282.83	0.0122	0.0172	-0.0173	-0.0050								
107	9/2/2008	75.47	-0.0293	1,166.36	-0.0908	0.0113	-0.0406	-0.1021								
108	10/1/2008	72.03	-0.0456	968.75	-0.1694	0.0067	-0.0523	-0.1761								
109	11/3/2008	78.34	0.0876	896.24	-0.0748	0.0019	0.0857	-0.0767								
110	12/1/2008	78.03	-0.0040	903.25	0.0078	0.0003	-0.0043	0.0075								
111	1/2/2009	74.75	-0.0420	825.88	-0.0857	0.0013	-0.0433	-0.0870								
112	2/2/2009	66.70	-0.1077	735.09	-0.1099	0.0030	-0.1107	-0.1129								
113	3/2/2009	66.90	0.0030	797.87	0.0854	0.0021	0.0009	0.0833								
114	4/1/2009	65.49	-0.0211	872.81	0.0939	0.0016	-0.0227	0.0923								
115	5/1/2009	68.53	0.0464	919.14	0.0531	0.0018	0.0446	0.0513								
116	6/1/2009	69.08	0.0080	919.32	0.0002	0.0018	0.0062	-0.0016								
117	7/1/2009	69.56	0.0069	987.48	0.0741	0.0018	0.0051	0.0723								
118	8/3/2009	68.75	-0.0116	1,020.62	0.0336	0.0017	-0.0133	0.0319								
119	9/1/2009	68.21	-0.0079	1,057.08	0.0357	0.0012	-0.0091	0.0345								
120	10/1/2009	71.26	0.0447	1,036.19	-0.0198	0.0007	0.0440	-0.0205								
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118	8/3/2009	68.75	-0.0116	1,020.62	0.0336	0.0017	-0.0133	0.0319									
119	9/1/2009	68.21	-0.0079	1,057.08	0.0357	0.0012	-0.0091	0.0345									
120	10/1/2009	71.26	0.0447	1,036.19	-0.0198	0.0007	0.0440	-0.0205									
121	11/2/2009	75.07	0.0535	1,095.63	0.0574	0.0005	0.0530	0.0569									
122	12/1/2009	68.19	-0.0916	1,115.10	0.0178	0.0005	-0.0921	0.0173									
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Figure 6 shows the results.

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12 Re	egression	1.000	0.070	0.070	29.096	0.000							
13 Re	esidual	117.000	0.282	0.002									
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1/ IN	Variable 1	-0.000	0.005	-1.1/3	0.243	-0.016	0.004	+ -0.010	0.004	+			
10 1	validule 1	0.484	0.090	3.594	0.000	0.300	0.002	0.300	0.00	<u>_</u>			
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Figure 6

We expect α to be zero. If equation (10) is true, then equation (11) may only be true if α is zero. If the estimated α is positive then XOM generates returns greater than what the CAPM predicts and should be a security that is bought for this reason (i.e. the security generates more return than it theoretically should). If the estimated α is negative then XOM generates returns less than what the CAPM predicts and should be sold if already owned. Our estimate for α is -0.006 and the p-value is 0.243 indicating that α is not significantly different from zero.

We expect β to be different from zero. If it is, then returns on the S&P 500 are useful in explaining returns on XOM. Our estimate for β is 0.484 and the p-value is 0.00000036. (Note Figure 6 shows the rounded result 0.000). This p-value is less than 0.01 so we may say that $R_{S\&P500}$ is significant at the 1% level.

The r-square is 0.199. So only 19.9% of the variation of XOM returns are explained by returns on the S&P500. Can we improve the model? Are there other factors that are significant and will explain more of the variation of XOM returns? We leave that question for the reader/class to answer.

Conclusion

By using basketball data, one can introduce regression in the classroom in a very intuitive manner. To get the full benefit from the example, the instructor needs to have the students suggest and debate what variables matter when modeling a player's points per game (PPG). The variable, minutes per game (MPG), becomes a very reasonable suggestion because if you do not play, you cannot score. In this manner, students begin to see the value of regression analysis and what types of questions can be answered using regression analysis.

The earlier portion of the paper can be introduced before the basketball data to demonstrate what a regression actually is (i.e. an estimation tool based on minimizing mean squared error) or can be introduced after the basketball data to see how a regression determines the structure (i.e. the coefficient estimates) of the model given a set of data.

We encourage the instructor to go through this portion of the exercise because it takes regression out of the computer and into a form where the student can logic through the reasoning behind using one set of coefficient estimates versus another set of coefficient estimates. Logically, the regression should have errors that average to zero and that the size of the error should be minimized in some manner as well. The depth of the discussion can vary, but the point of the discussion should be that some coefficient estimates are better than others and that an optimal set of coefficient estimates should be found.

After working through both the first portion of the paper and the basketball portion of the paper, applying economic or financial data becomes the next logical step. We have presented a fairly easy example for the Capital Asset Pricing Model, but many other examples could be performed as well. We suggest using readily available data from the Internet or supplied by the instructor because the overall lesson should be focused on regression and not data issues.

After the students are more comfortable with regression, data issues can be introduced. One way to segue into such a conversation is to ask, what kind of basketball data would be better for modeling PPG? There is a short discussion earlier in the basketball section that considers a home-away variable and data frequency, however, let the students work through this issue through discussion because developing their intuition matters the most.

References:

Sharpe, William. 1964. "Capital asset prices: a theory of market equilibrium under conditions of risk." *Journal of Finance*, 19 (3): 425-442.